

# Dynamic Linear Models

FISH 550 – Applied Time Series Analysis

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# Topics for today

## Univariate response

- Stochastic level & growth
- Dynamic Regression
- Dynamic Regression with fixed season
- Forecasting with a DLM
- Model diagnostics

## Multivariate response

# Code for today

You can find the R code for these lecture notes and other related exercises [here](#).

# Simple linear regression

Let's begin with a linear regression model

$$y_i = \alpha + \beta x_i + e_i \text{ with } e_i \sim N(0, \sigma^2)$$

The index  $i$  has no explicit meaning in that shuffling  $(y_i, x_i)$  pairs has no effect on parameter estimation

# Simple linear regression

We can write the model in matrix form

# Simple linear regression

We can write the model in matrix form

with

$$\mathbf{X}_i^\top = [1 \quad x_i] \text{ and } \boldsymbol{\theta} = [\alpha \quad \beta]^\top$$

# Dynamic linear model (DLM)

In a *dynamic* linear model, the regression parameters change over time, so we write

$$y_i = \mathbf{X}_i^\top \boldsymbol{\theta} + e_i \quad (\text{static})$$

as

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t \quad (\text{dynamic})$$

# Dynamic linear model (DLM)

There are 2 important points here:

$$y_{\boxed{t}} = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t$$

1. Subscript  $t$  explicitly acknowledges implicit info in the time ordering of the data in  $\mathbf{y}$

# Dynamic linear model (DLM)

There are 2 important points here:

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_{\boxed{t}} + e_t$$

1. Subscript  $t$  explicitly acknowledges implicit info in the time ordering of the data in  $\mathbf{y}$
2. The relationship between  $\mathbf{y}$  and  $\mathbf{X}$  is unique for every  $t$

# Constraining a DLM

Close examination of the DLM reveals an apparent problem for parameter estimation

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t$$

# Constraining a DLM

Close examination of the DLM reveals an apparent problem for parameter estimation

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t$$

We only have 1 data point per time step (ie,  $y_t$  is a scalar)

Thus, we can only estimate 1 parameter (with no uncertainty)!

# Constraining a DLM

To address this issue, we'll constrain the regression parameters to be dependent from  $t$  to  $t + 1$

$$\boldsymbol{\theta}_t = \mathbf{G}_t \boldsymbol{\theta}_{t-1} + \mathbf{w}_t \text{ with } \mathbf{w}_t \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

# Constraining a DLM

In practice, we often make  $\mathbf{G}_t$  time invariant

$$\boldsymbol{\theta}_t = \mathbf{G}\boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

# Constraining a DLM

In practice, we often make  $\mathbf{G}_t$  time invariant

$$\boldsymbol{\theta}_t = \mathbf{G}\boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

or assume  $\mathbf{G}_t$  is an  $m \times m$  identity matrix  $\mathbf{I}_m$

$$\begin{aligned}\boldsymbol{\theta}_t &= \mathbf{I}_m\boldsymbol{\theta}_{t-1} + \mathbf{w}_t \\ &= \boldsymbol{\theta}_{t-1} + \mathbf{w}_t\end{aligned}$$

In the latter case, the parameters follow a random walk over time

# DLM in state-space form

Observation model relates the covariates  $\mathbf{X}$  to the data

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t$$

State model determines how parameters “evolve” over time

$$\boldsymbol{\theta}_t = \mathbf{G}\boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

# DLM in MARSS notation

Full state-space form

where

$$\mathbf{Z}_t = \mathbf{X}_t^\top, \mathbf{x}_t = \boldsymbol{\theta}_t, v_t = e_t, \mathbf{B} = \mathbf{G}$$

# Contrast in covariate effects

Note: DLMs include covariate effect in the observation eqn much differently than other forms of MARSS models

DLM:  $\mathbf{Z}_t$  is covariates,  $\mathbf{x}_t$  is parameters

$$y_t = \boxed{\mathbf{Z}_t \mathbf{x}_t} + v_t$$

Others:  $\mathbf{d}_t$  is covariates,  $\mathbf{D}$  is parameters

$$y_t = \mathbf{Z}_t \mathbf{x}_t + \boxed{\mathbf{D} \mathbf{d}_t} + v_t$$

# Other forms of DLMs

The regression model is but one type

Others include:

- stochastic “level” (intercept)
- stochastic “growth” (trend, bias)
- seasonal effects (fixed, harmonic)

# The most simple DLM

Stochastic level

$$\begin{aligned}y_t &= \alpha_t + e_t \\ \alpha_t &= \alpha_{t-1} + w_t\end{aligned}$$

# The most simple DLM

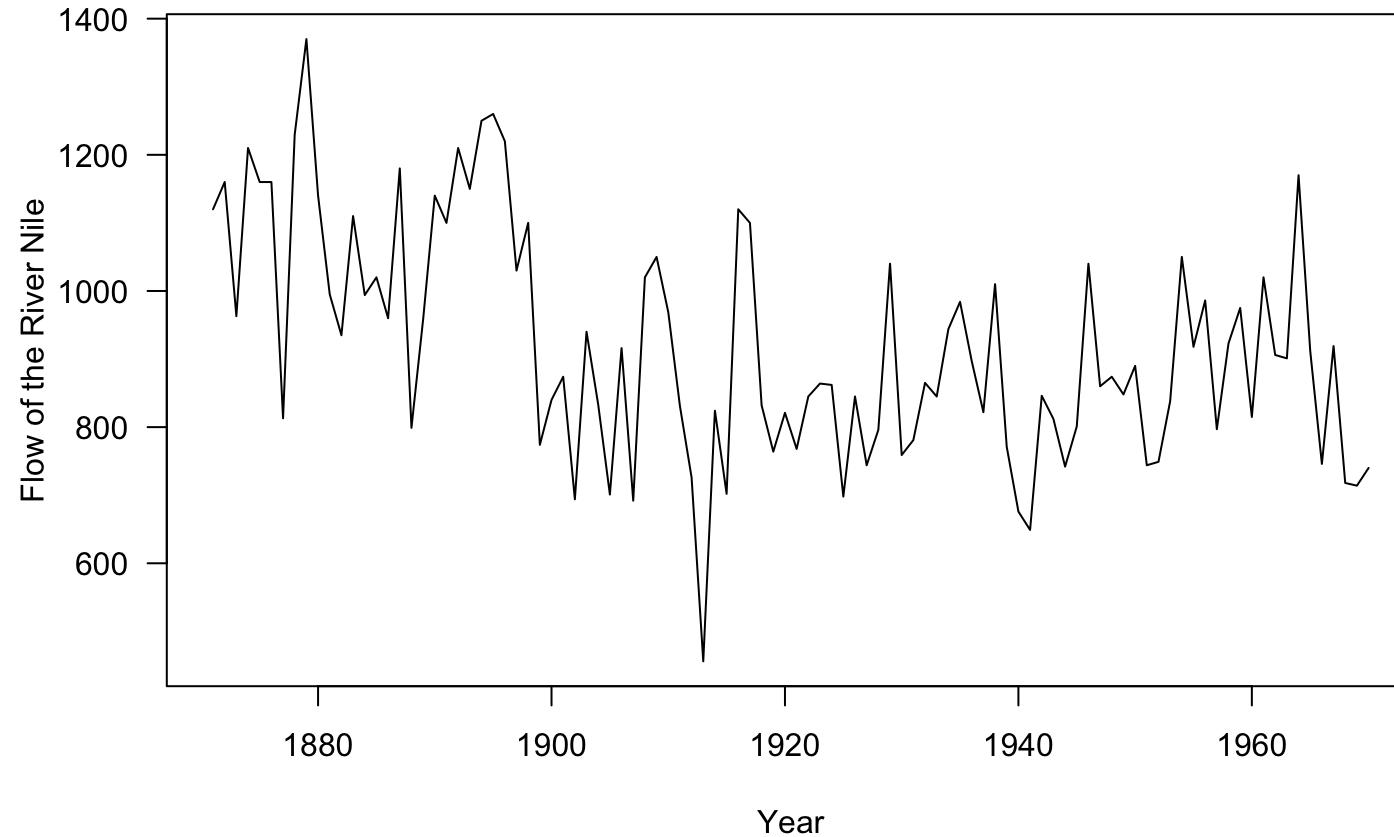
Stochastic level = random walk with obs error

$$\begin{aligned}y_t &= \alpha_t + e_t \\ \alpha_t &= \alpha_{t-1} + w_t\end{aligned}$$

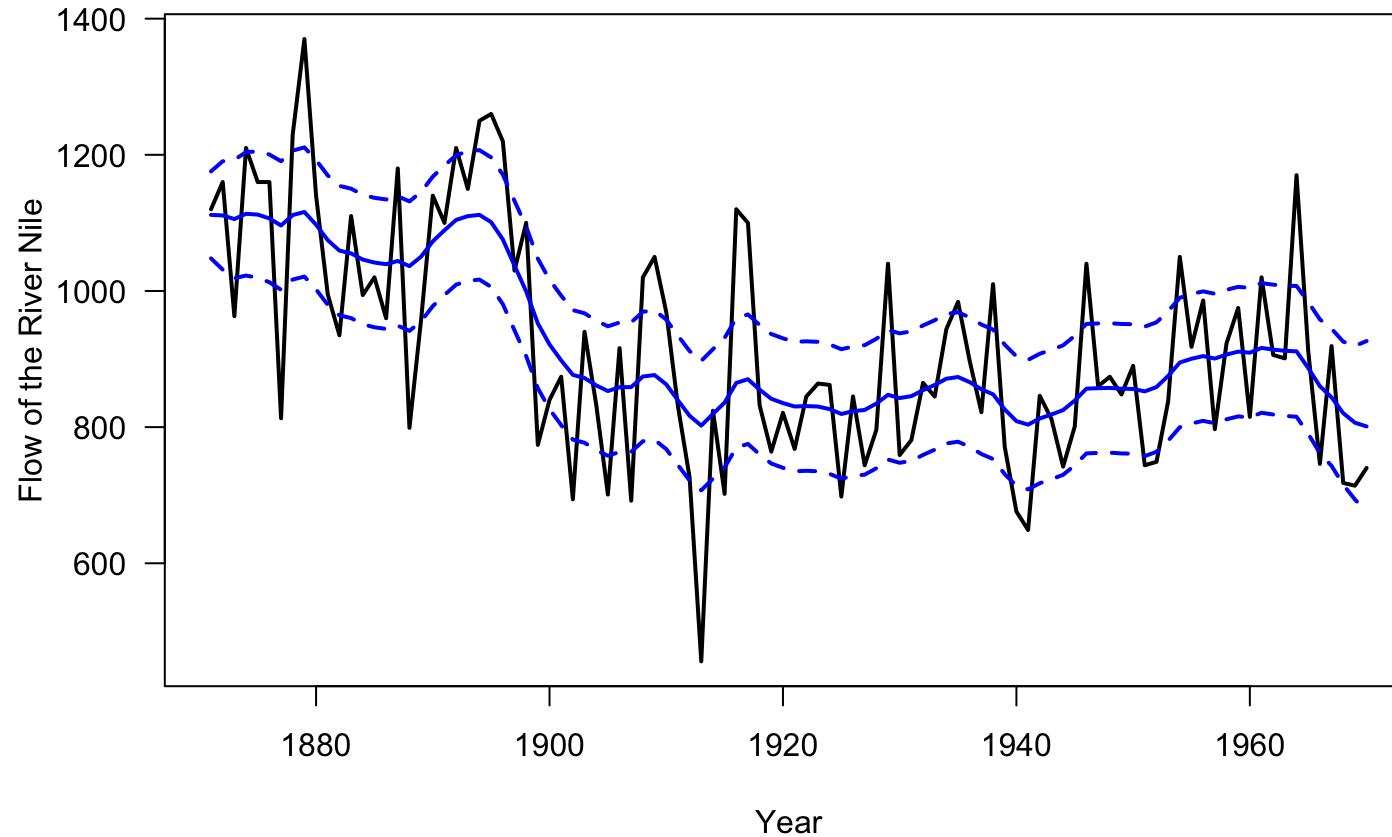


$$\begin{aligned}y_t &= x_t + v_t \\ x_t &= x_{t-1} + w_t\end{aligned}$$

# Example of stochastic level model



# Example of stochastic level model



# Univariate DLM for level & growth

Stochastic “level”  $\alpha_t$  with deterministic “growth”  $\eta$

$$y_t = \alpha_t + e_t$$
$$\alpha_t = \alpha_{t-1} + \eta + w_t$$

# Univariate DLM for level & growth

Stochastic “level”  $\alpha_t$  with deterministic “growth”  $\eta$

$$\begin{aligned}y_t &= \alpha_t + e_t \\ \alpha_t &= \alpha_{t-1} + \eta + w_t \\ &\Downarrow \\ y_t &= x_t + v_t \\ x_t &= x_{t-1} + u + w_t\end{aligned}$$

This is just a random walk with bias  $u$

# Univariate DLM for level & growth

Stochastic “level”  $\alpha_t$  with stochastic “growth”  $\eta_t$

$$\begin{aligned}y_t &= \alpha_t + e_t \\ \alpha_t &= \alpha_{t-1} + \eta_{t-1} + w_{\alpha,t} \\ \eta_t &= \eta_{t-1} + w_{\eta,t}\end{aligned}$$

Now the “growth” term  $\eta_t$  evolves as well

# Univariate DLM for level & growth

Evolution of  $\alpha_t$  and  $\eta_t$

$$\begin{aligned}\alpha_t &= \alpha_{t-1} + \eta_{t-1} + w_{\alpha,t} \\ \eta_t &= \eta_{t-1} + w_{\eta,t}\end{aligned}$$

How do we make this work in practice?

# Univariate DLM for level & growth

Evolution of  $\alpha_t$  and  $\eta_t$

$$\alpha_t = \alpha_{t-1} + \eta_{t-1} + w_{\alpha,t}$$

$$\eta_t = \eta_{t-1} + w_{\eta,t}$$

$\Downarrow$

$$\alpha_t = 1\alpha_{t-1} + 1\eta_{t-1} + w_{\alpha,t}$$

$$\eta_t = 0\alpha_{t-1} + 1\eta_{t-1} + w_{\eta,t}$$

Rewrite the equations with explicit coefficients on  $\alpha_{t-1}$  and  $\eta_{t-1}$

# Univariate DLM for level & growth

Evolution of  $\alpha_t$  and  $\eta_t$

$$\alpha_t = \alpha_{t-1} + \eta_{t-1} + w_{\alpha,t}$$

$$\eta_t = \eta_{t-1} + w_{\eta,t}$$

$\Downarrow$

$$\alpha_t = \underline{1}\alpha_{t-1} + \underline{1}\eta_{t-1} + w_{\alpha,t}$$

$$\eta_t = \underline{0}\alpha_{t-1} + \underline{1}\eta_{t-1} + w_{\eta,t}$$

$\Downarrow$

$$\underbrace{\begin{bmatrix} \alpha_t \\ \eta_t \end{bmatrix}}_{\theta_t} = \underbrace{\begin{bmatrix} \underline{1} & \underline{1} \\ \underline{0} & \underline{1} \end{bmatrix}}_{\mathbf{G}} \underbrace{\begin{bmatrix} \alpha_{t-1} \\ \eta_{t-1} \end{bmatrix}}_{\theta_{t-1}} + \underbrace{\begin{bmatrix} w_{\alpha,t} \\ w_{\eta,t} \end{bmatrix}}_{\mathbf{w}_t}$$

# Univariate DLM for level & growth

Evolution of  $\alpha_t$  and  $\eta_t$  in MARSS form

$$x_{1,t} = x_{1,t-1} + x_{2,t-1} + w_{1,t}$$

$$x_{2,t} = x_{2,t-1} + w_{2,t}$$

$\Downarrow$

$$x_{1,t} = \underline{1}x_{1,t-1} + \underline{1}x_{2,t-1} + w_{1,t}$$

$$x_{2,t} = \underline{0}x_{1,t-1} + \underline{1}x_{2,t-1} + w_{2,t}$$

$\Downarrow$

$$\underbrace{\begin{bmatrix} x_{1,t} \\ x_{2,t} \end{bmatrix}}_{\mathbf{x}_t} = \underbrace{\begin{bmatrix} \underline{1} & \underline{1} \\ \underline{0} & \underline{1} \end{bmatrix}}_{\mathbf{B}} \underbrace{\begin{bmatrix} x_{1,t-1} \\ x_{2,t-1} \end{bmatrix}}_{\mathbf{x}_{t-1}} + \underbrace{\begin{bmatrix} w_{1,t} \\ w_{2,t} \end{bmatrix}}_{\mathbf{w}_t}$$

# Univariate DLM for level & growth

Observation model for stochastic *level & growth*

$$\begin{aligned} y_t &= \alpha_t + \nu_t \\ &\Downarrow \\ y_t &= \underline{1}\alpha_t + \underline{0}\eta_t + \nu_t \end{aligned}$$

Again, rewrite equation with explicit coefficients on  $\alpha_t$  and  $\eta_t$

# Univariate DLM for level & growth

Observation model for stochastic *level & growth*

$$y_t = \alpha_t + \nu_t$$



$$y_t = \underline{1}\alpha_t + \underline{0}\eta_t + \nu_t$$



$$y_t = \underbrace{\begin{bmatrix} \underline{1} & \underline{0} \end{bmatrix}}_{\mathbf{X}_t^\top} \underbrace{\begin{bmatrix} \alpha_t \\ \eta_t \end{bmatrix}}_{\boldsymbol{\theta}_t} + \nu_t$$

# Univariate DLM for level & growth

Obs model for stochastic *level & growth* in MARSS form

$$y_t = x_t + \nu_t$$



$$y_t = \underline{1}x_{1,t} + \underline{0}x_{2,t} + \nu_t$$



$$y_t = \underbrace{\begin{bmatrix} \underline{1} & \underline{0} \end{bmatrix}}_{\mathbf{Z}_t} \underbrace{\begin{bmatrix} x_{1,t} \\ x_{2,t} \end{bmatrix}}_{\mathbf{x}_t} + \nu_t$$

# Univariate DLM for regression

Stochastic intercept and slope

$$y_t = \alpha_t + \beta_t x_t + \nu_t$$

# Univariate DLM for regression

Stochastic intercept and slope

$$\begin{aligned}y_t &= \alpha_t + \beta_t x_t + \nu_t \\&\Downarrow \\y_t &= \underline{1}\alpha_t + \underline{x_t}\beta_t + \nu_t\end{aligned}$$

Rewrite the equation with explicit coefficients for  $\alpha_t$  &  $\beta_t$

# Univariate DLM for regression

Stochastic intercept and slope

$$y_t = \alpha_t + \beta_t x_t + v_t$$



$$y_t = \underline{1} \alpha_t + \underline{x_t} \beta_t + v_t$$



$$y_t = \underbrace{\begin{bmatrix} 1 & x_t \end{bmatrix}}_{\mathbf{X}_t^\top} \underbrace{\begin{bmatrix} \alpha_t \\ \beta_t \end{bmatrix}}_{\boldsymbol{\theta}_t} + v_t$$

# Univariate DLM for regression

Stochastic intercept and slope in MARSS form

$$\begin{aligned} y_t &= x_{1,t} + x_{1,t}z_{2,t} + v_t \\ &\Downarrow \\ y_t &= \underline{1}x_{1,t} + \underline{z_{2,t}}x_{2,t} + v_t \\ &\Downarrow \\ y_t &= \underbrace{\begin{bmatrix} 1 & z_{2,t} \end{bmatrix}}_{\mathbf{Z}_t} \underbrace{\begin{bmatrix} x_{1,t} \\ x_{2,t} \end{bmatrix}}_{\mathbf{x}_t} + v_t \end{aligned}$$

# Univariate DLM for regression

Parameter evolution follows a random walk

$$\begin{aligned}\alpha_t &= \alpha_{t-1} + w_{\alpha,t} \\ \beta_t &= \beta_{t-1} + w_{\beta,t} \\ &\Downarrow \\ \underbrace{\begin{bmatrix} \alpha_t \\ \beta_t \end{bmatrix}}_{\theta_t} &= \underbrace{\begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \end{bmatrix}}_{\theta_{t-1}} + \underbrace{\begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \end{bmatrix}}_{\mathbf{w}_t}\end{aligned}$$

# Univariate DLM for regression

Parameter evolution in MARSS form

$$\begin{aligned}x_{1,t} &= x_{1,t-1} + w_{1,t} \\x_{2,t} &= x_{2,t-1} + w_{2,t} \\&\Downarrow \\ \underbrace{\begin{bmatrix} x_{1,t} \\ x_{2,t} \end{bmatrix}}_{\mathbf{x}_t} &= \underbrace{\begin{bmatrix} x_{1,t-1} \\ x_{2,t-1} \end{bmatrix}}_{\mathbf{x}_{t-1}} + \underbrace{\begin{bmatrix} w_{1,t} \\ w_{2,t} \end{bmatrix}}_{\mathbf{w}_t}\end{aligned}$$

# Univariate DLM with seasonal effect

Dynamic linear regression with fixed seasonal effect

$$y_t = \alpha_t + \beta_t x_t + \gamma_{qtr} + e_t$$

$$\gamma_{qtr} = \begin{cases} \gamma_1 & \text{if } qtr = 1 \\ \gamma_2 & \text{if } qtr = 2 \\ \gamma_3 & \text{if } qtr = 3 \\ \gamma_4 & \text{if } qtr = 4 \end{cases}$$

# Univariate DLM with seasonal effect

Dynamic linear regression with fixed seasonal effect

$$y_t = \alpha_t + \beta_t x_t + \gamma_{qtr} + e_t$$



$$y_t = \begin{bmatrix} 1 & x_t & 1 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_{qtr} \end{bmatrix} + e_t$$

Rewrite the equation with explicit coefficients on parameters

# Univariate DLM with seasonal effect

Evolution of parameters

$$\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_{qtr} \end{bmatrix} = \begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ ? \end{bmatrix} + \begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ ? \end{bmatrix}$$

How should we model the fixed effect of  $\gamma_{qtr}$ ?

# Univariate DLM with seasonal effect

Evolution of parameters

$$\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_{qtr} \end{bmatrix} = \begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ \gamma_{qtr} \end{bmatrix} + \begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ 0 \end{bmatrix}$$

We don't want  $\gamma_{qtr}$  to evolve as a function of the previous  $t$

# Univariate DLM with seasonal effect

Evolution of parameters

$$\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_{qtr} \end{bmatrix} = \begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ \gamma_{qtr} \end{bmatrix} + \begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ 0 \end{bmatrix}$$

OK, so how do we select the right quarterly effect?

# Univariate DLM with seasonal effect

Separate out the quarterly effects

$$y_t = \alpha_t + \beta_t x_t + \gamma_{qtr} + e_t$$



$$y_t = \alpha_t + \beta_t x_t + \gamma_1 + \gamma_2 + \gamma_3 + \gamma_4 + e_t$$

# Univariate DLM with seasonal effect

Rewrite quarterly effects in matrix notation

$$y_t = \alpha_t + \beta_t x_t + \gamma_1 + \gamma_2 + \gamma_3 + \gamma_4 + e_t$$

↓

$$y_t = \begin{bmatrix} 1 & x_t & 1 & 1 & 1 & 1 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}$$

But how do we select only the current quarter?

# Univariate DLM with seasonal effect

We could set some values in  $\mathbf{x}_t$  to 0 ( $qtr = 1$ )

$$y_t = \begin{bmatrix} 1 & x_t & 1 & 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}$$

$\Downarrow$

$$y_t = \alpha_t + \beta_t x_t + \gamma_1 + e_t$$

# Univariate DLM with seasonal effect

We could set some values in  $\mathbf{x}_t$  to 0 ( $qtr = 2$ )

$$y_t = \begin{bmatrix} 1 & x_t & 0 & 1 & 0 & 0 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}$$

$\Downarrow$

$$y_t = \alpha_t + \beta_t x_t + \gamma_2 + e_t$$

# Univariate DLM with seasonal effect

But *how* would we set the correct 0/1 values?

$$\mathbf{X}_t^\top = [1 \quad x_t \quad ? \quad ? \quad ? \quad ?]$$

# Univariate DLM with seasonal effect

We could instead reorder the  $\gamma_i$  within  $\theta_t$  ( $qtr = 1$ )

$$y_t = \begin{bmatrix} 1 & x_t & 1 & 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}$$

$\Downarrow$

$$y_t = \alpha_t + \beta_t x_t + \gamma_1 + e_t$$

# Univariate DLM with seasonal effect

We could instead reorder the  $\gamma_i$  within  $\theta_t$  ( $qtr = 2$ )

$$y_t = \begin{bmatrix} 1 & x_t & 1 & 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \\ \gamma_1 \end{bmatrix}$$

$\Downarrow$

$$y_t = \alpha_t + \beta_t x_t + \gamma_2 + e_t$$

# Univariate DLM with seasonal effect

But *how* would we shift the  $\gamma_i$  within  $\theta_t$ ?

$$\theta_t = \begin{bmatrix} \alpha_t \\ \beta_t \\ ? \\ ? \\ ? \\ ? \end{bmatrix}$$

# Example of non-diagonal $\mathbf{G}$

We can use a non-diagonal submatrix in the lower right of  $\mathbf{G}$  to get the correct quarter effect

$$\mathbf{G} = \left[ \begin{array}{cc|cccc} 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ \hline 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 1 & 0 & 0 & 0 \end{array} \right]$$

# Evolving parameters

Quarter 1

$$\underbrace{\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}}_{\theta_t} = \underbrace{\begin{bmatrix} 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 1 & 0 & 0 & 0 \end{bmatrix}}_{\mathbf{G}} \underbrace{\begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ \gamma_4 \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \end{bmatrix}}_{\theta_{t-1}} + \underbrace{\begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}}_{\mathbf{w}_t}$$

# Evolving parameters

Quarter 2

$$\underbrace{\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \\ \gamma_1 \end{bmatrix}}_{\theta_t} = \underbrace{\begin{bmatrix} 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 1 & 0 & 0 & 0 \end{bmatrix}}_{\mathbf{G}} \underbrace{\begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}}_{\theta_{t-1}} + \underbrace{\begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}}_{\mathbf{w}_t}$$

# Evolving parameters

Quarter 3

$$\underbrace{\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_3 \\ \gamma_4 \\ \gamma_1 \\ \gamma_2 \end{bmatrix}}_{\theta_t} = \underbrace{\begin{bmatrix} 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 1 & 0 & 0 & 0 \end{bmatrix}}_{\mathbf{G}} \underbrace{\begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \\ \gamma_1 \end{bmatrix}}_{\theta_{t-1}} + \underbrace{\begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}}_{\mathbf{w}_t}$$

# Forecasting with a DLM

# Forecasting with a DLM

DLMs are often used in a forecasting context where we want a prediction for time  $t$  based on the data up through time  $t - 1$

# Forecasting with a DLM

## Pseudo-code

1. get estimate of today's parameters from yesterday's
2. make prediction based on today's parameters & covariates
3. get observation for today
4. update parameters and repeat

# Forecasting with a DLM

Step 1: Define the parameters at time  $t = 0$

$$\theta_0 | y_0 = \pi + w_1 \text{ with } w_1 \sim \text{MVN}(\mathbf{0}, \Lambda)$$

# Forecasting with a DLM

Step 1: Define the parameters at time  $t = 0$

$$\theta_0 | y_0 = \pi + w_1 \text{ with } w_1 \sim \text{MVN}(\mathbf{0}, \Lambda)$$



$$E(\theta_0) = \pi$$

# Forecasting with a DLM

Step 1: Define the parameters at time  $t = 0$

$$\boldsymbol{\theta}_0 | y_0 = \boldsymbol{\pi} + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \boldsymbol{\Lambda})$$



$$E(\boldsymbol{\theta}_0) = \boldsymbol{\pi}$$

and

$$\text{Var}(\boldsymbol{\theta}_0) = \text{Var}(\boldsymbol{\pi}) + \text{Var}(\mathbf{w}_1)$$

$$\text{Var}(\boldsymbol{\theta}_0) = \mathbf{0} + \boldsymbol{\Lambda}$$

$$\text{Var}(\boldsymbol{\theta}_0) = \boldsymbol{\Lambda}$$

# Forecasting with a DLM

Step 1: Define the parameters at time  $t = 0$

$$\boldsymbol{\theta}_0 | y_0 = \boldsymbol{\pi} + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \boldsymbol{\Lambda})$$



$$E(\boldsymbol{\theta}_0) = \boldsymbol{\pi}$$

and

$$\text{Var}(\boldsymbol{\theta}_0) = \boldsymbol{\Lambda}$$



$$\boldsymbol{\theta}_0 | y_0 \sim \text{MVN}(\boldsymbol{\pi}, \boldsymbol{\Lambda})$$

# Forecasting with a DLM

Step 2: Define the parameters at time  $t = 1$

$$\boldsymbol{\theta}_1 | y_0 = \mathbf{G}\boldsymbol{\theta}_0 + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

# Forecasting with a DLM

Step 2: Define the parameters at time  $t = 1$

$$\boldsymbol{\theta}_1 | y_0 = \mathbf{G}\boldsymbol{\theta}_0 + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$



$$E(\boldsymbol{\theta}_1) = \mathbf{G}E(\boldsymbol{\theta}_0)$$

$$E(\boldsymbol{\theta}_1) = \mathbf{G}\boldsymbol{\pi}$$

# Forecasting with a DLM

Step 2: Define the parameters at time  $t = 1$

$$\boldsymbol{\theta}_1 | y_0 = \mathbf{G}\boldsymbol{\theta}_0 + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$



$$E(\boldsymbol{\theta}_1) = \mathbf{G}\boldsymbol{\pi}$$

and

$$\text{Var}(\boldsymbol{\theta}_1) = \mathbf{G}\text{Var}(\boldsymbol{\theta}_0)\mathbf{G}^\top + \text{Var}(\mathbf{w}_1)$$

$$\text{Var}(\boldsymbol{\theta}_1) = \mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q}$$

# Forecasting with a DLM

Step 2: Define the parameters at time  $t = 1$

$$\boldsymbol{\theta}_1 | y_0 = \mathbf{G}\boldsymbol{\theta}_0 + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

$$\Downarrow$$

$$E(\boldsymbol{\theta}_1) = \mathbf{G}\boldsymbol{\pi}$$

and

$$\text{Var}(\boldsymbol{\theta}_1) = \mathbf{G}\Lambda\mathbf{G}^\top + \mathbf{Q}$$

$$\Downarrow$$

$$\boldsymbol{\theta}_1 | y_0 \sim \text{MVN}(\mathbf{G}\boldsymbol{\pi}, \mathbf{G}\Lambda\mathbf{G}^\top + \mathbf{Q})$$

# Forecasting with a DLM

Step 3: Make a forecast of  $y_t$  at time  $t = 1$

$$y_1 | y_0 = \mathbf{X}_1^\top \boldsymbol{\theta}_1 + e_1 \text{ with } e_1 \sim N(0, R)$$



$$E(y_1) = \mathbf{X}_1^\top E(\boldsymbol{\theta}_1)$$

$$E(y_1) = \mathbf{X}_1^\top \mathbf{G}\boldsymbol{\pi}$$

# Forecasting with a DLM

Step 3: Make a forecast of  $y_t$  at time  $t = 1$

$$y_1 | y_0 = \mathbf{X}_1^\top \boldsymbol{\theta}_1 + e_1 \text{ with } e_1 \sim N(0, R)$$



$$E(y_1) = \mathbf{X}_1^\top \mathbf{G}\boldsymbol{\pi}$$

and

$$\text{Var}(y_1) = \mathbf{X}_1^\top \text{Var}(\boldsymbol{\theta}_1) \mathbf{X}_1 + \text{Var}(e_1)$$

$$\text{Var}(y_1) = \mathbf{X}_1^\top [\mathbf{G}\Lambda\mathbf{G}^\top + \mathbf{Q}] \mathbf{X}_1 + R$$

# Forecasting with a DLM

Step 3: Make a forecast of  $y_t$  at time  $t = 1$

$$y_1 | y_0 = \mathbf{X}_1^\top \boldsymbol{\theta}_1 + e_1 \text{ with } e_1 \sim N(0, R)$$



$$E(y_1) = \mathbf{X}_1^\top \mathbf{G}\boldsymbol{\pi}$$

and

$$\text{Var}(y_1) = \mathbf{X}_1^\top [\mathbf{G}\Lambda\mathbf{G}^\top + \mathbf{Q}] \mathbf{X}_1 + R$$



$$y_1 | y_0 \sim N(\mathbf{X}_1^\top [\mathbf{G}\boldsymbol{\pi}], \mathbf{X}_1^\top [\mathbf{G}\Lambda\mathbf{G}^\top + \mathbf{Q}] \mathbf{X}_1 + R)$$

# Forecasting with a DLM

Putting it all together

$$\boldsymbol{\theta}_0 | y_0 \sim \text{MVN}(\boldsymbol{\pi}, \boldsymbol{\Lambda})$$

$$\boldsymbol{\theta}_t | y_{t-1} \sim \text{MVN}(\mathbf{G}\boldsymbol{\pi}, \mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q})$$

$$y_t | y_{t-1} \sim \text{N}(\mathbf{X}_t^\top [\mathbf{G}\boldsymbol{\pi}], \mathbf{X}_t^\top [\mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q}] \mathbf{X}_t + R)$$

# Forecasting with a DLM

Putting it all together

$$\boldsymbol{\theta}_0 | y_0 \sim \text{MVN}(\boldsymbol{\pi}, \boldsymbol{\Lambda})$$

$$\boldsymbol{\theta}_t | y_{t-1} \sim \text{MVN}(\mathbf{G}\boldsymbol{\pi}, \mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q})$$

$$y_t | y_{t-1} \sim \text{N}(\mathbf{X}_t^\top [\mathbf{G}\boldsymbol{\pi}], \mathbf{X}_t^\top [\mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q}] \mathbf{X}_t + R)$$

Using `MARSS( )` will make this easy to do

# Diagnostics for DLMs

Just as with other models, we'd like to know if our fitted DLM meets its underlying assumptions

We can calculate the forecast error  $e_t$  as

$$e_t = y_t - \hat{y}_t$$

and check if

- (1)  $e_t \sim N(0, \sigma)$
- (2)  $\text{Cov}(e_t, e_{t-1}) = 0$

with a QQ-plot (1) and an ACF (2)

# Multivariate DLMs

# Multivariate DLMs

We can expand our DLM to have a multivariate response

# The most simple multivariate DLM

Multiple observations of a stochastic level

$$\begin{aligned}\mathbf{y}_t &= \mathbf{Z}\boldsymbol{\alpha}_t + \mathbf{v}_t & \mathbf{y}_t \text{ is } n \times 1 \\ \boldsymbol{\alpha}_t &= \boldsymbol{\alpha}_{t-1} + w_t & \boldsymbol{\alpha}_t \text{ is } 1 \times 1\end{aligned}$$

with

$$\mathbf{Z} = \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix}$$

# The most simple multivariate DLM

Multiple observations of a random walk

$$\begin{aligned}\mathbf{y}_t &= \mathbf{Z}x_t + \mathbf{v}_t & \mathbf{y}_t \text{ is } n \times 1 \\ x_t &= x_{t-1} + w_t & x_t \text{ is } 1 \times 1\end{aligned}$$

with

$$\mathbf{Z} = \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix}$$

# Another simple multivariate DLM

Multiple observations of multiple levels

$$\begin{aligned}\mathbf{y}_t &= \mathbf{Z}\boldsymbol{\alpha}_t + \mathbf{v}_t & \mathbf{y}_t \text{ is } n \times 1 \\ \boldsymbol{\alpha}_t &= \boldsymbol{\alpha}_{t-1} + \mathbf{w}_t & \boldsymbol{\alpha}_t \text{ is } n \times 1\end{aligned}$$

with

$$\mathbf{Z} = \mathbf{I}_n = \begin{bmatrix} 1 & 0 & \dots & 0 \\ 0 & 1 & \ddots & 0 \\ \vdots & \ddots & \ddots & 0 \\ 0 & \dots & 0 & 1 \end{bmatrix}$$

# Another simple multivariate DLM

Multiple observations of multiple random walks

$$\begin{aligned}\mathbf{y}_t &= \mathbf{Z}\mathbf{x}_t + \mathbf{v}_t & \mathbf{y}_t \text{ is } n \times 1 \\ \mathbf{x}_t &= \mathbf{x}_{t-1} + \mathbf{w}_t & \mathbf{x}_t \text{ is } n \times 1\end{aligned}$$

with

$$\mathbf{Z} = \mathbf{I}_n = \begin{bmatrix} 1 & 0 & \dots & 0 \\ 0 & 1 & \ddots & 0 \\ \vdots & \ddots & \ddots & 0 \\ 0 & \dots & 0 & 1 \end{bmatrix}$$

# Multivariate DLMs

Regression model

Our univariate model

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t \text{ with } e_t \sim \mathcal{N}(0, R)$$

becomes

$$\mathbf{y}_t = (\mathbf{X}_t^\top \otimes \mathbf{I}_n) \boldsymbol{\theta}_t + \mathbf{e}_t \text{ with } \mathbf{e}_t \sim \text{MVN}(\mathbf{0}, \mathbf{R})$$

where  $\otimes$  is the *Kronecker product*

# Kronecker products

If  $\mathbf{A}$  is an  $m \times n$  matrix and  $\mathbf{B}$  is a  $p \times q$  matrix

then  $\mathbf{A} \otimes \mathbf{B}$  will be an  $mp \times nq$  matrix

$$\mathbf{A} \otimes \mathbf{B} = \begin{bmatrix} a_{11}\mathbf{B} & \dots & a_{1n}\mathbf{B} \\ \vdots & \ddots & \vdots \\ a_{m1}\mathbf{B} & \dots & a_{mn}\mathbf{B} \end{bmatrix}$$

# Kronecker products

For example

$$\mathbf{A} = \begin{bmatrix} 1 & 2 \\ 3 & 4 \end{bmatrix} \text{ and } \mathbf{B} = \begin{bmatrix} 2 & 4 \\ 6 & 8 \end{bmatrix}$$

so

$$\mathbf{A} \otimes \mathbf{B} = \begin{bmatrix} 1 \begin{bmatrix} 2 & 4 \\ 6 & 8 \end{bmatrix} & 2 \begin{bmatrix} 2 & 4 \\ 6 & 8 \end{bmatrix} \\ 3 \begin{bmatrix} 2 & 4 \\ 6 & 8 \end{bmatrix} & 4 \begin{bmatrix} 2 & 4 \\ 6 & 8 \end{bmatrix} \end{bmatrix} = \begin{bmatrix} 2 & 4 & 4 & 8 \\ 6 & 8 & 12 & 16 \\ 6 & 12 & 8 & 16 \\ 18 & 24 & 24 & 32 \end{bmatrix}$$

# Multivariate DLMs

Regression model with  $n = 2$

$$\mathbf{y}_t = (\mathbf{X}_t^\top \otimes \mathbf{I}_n) \boldsymbol{\theta}_t + \mathbf{e}_t$$



$$\begin{bmatrix} y_{1,t} \\ y_{2,t} \end{bmatrix} = \left( \begin{bmatrix} 1 & x_t \end{bmatrix} \otimes \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \right) \begin{bmatrix} \alpha_{1,t} \\ \alpha_{2,t} \\ \beta_{1,t} \\ \beta_{2,t} \end{bmatrix} + \begin{bmatrix} e_{1,t} \\ e_{2,t} \end{bmatrix}$$

# Multivariate DLMs

$$\begin{bmatrix} y_{1,t} \\ y_{2,t} \end{bmatrix} = \left( \begin{bmatrix} 1 & x_t \end{bmatrix} \otimes \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \right) \begin{bmatrix} \alpha_{1,t} \\ \alpha_{2,t} \\ \beta_{1,t} \\ \beta_{2,t} \end{bmatrix} + \begin{bmatrix} e_{1,t} \\ e_{2,t} \end{bmatrix}$$

↓

$$\begin{bmatrix} y_{1,t} \\ y_{2,t} \end{bmatrix} = \begin{bmatrix} 1 & 0 & x_t & 0 \\ 0 & 1 & 0 & x_t \end{bmatrix} \begin{bmatrix} \alpha_{1,t} \\ \alpha_{2,t} \\ \beta_{1,t} \\ \beta_{2,t} \end{bmatrix} + \begin{bmatrix} e_{1,t} \\ e_{2,t} \end{bmatrix}$$

# Multivariate DLMs

Covariance of observation errors

$$\mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma & 0 & 0 & 0 \\ 0 & \sigma & 0 & 0 \\ 0 & 0 & \sigma & 0 \\ 0 & 0 & 0 & \sigma \end{bmatrix} \text{ or } \mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma_1 & 0 & 0 & 0 \\ 0 & \sigma_2 & 0 & 0 \\ 0 & 0 & \sigma_3 & 0 \\ 0 & 0 & 0 & \sigma_4 \end{bmatrix}$$

$$\mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma & \gamma & \gamma & \gamma \\ \gamma & \sigma & \gamma & \gamma \\ \gamma & \gamma & \sigma & \gamma \\ \gamma & \gamma & \gamma & \sigma \end{bmatrix} \text{ or } \mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma_1 & 0 & 0 & 0 \\ 0 & \sigma_2 & 0 & \gamma_{2,4} \\ 0 & 0 & \sigma_3 & 0 \\ 0 & \gamma_{2,4} & 0 & \sigma_4 \end{bmatrix}$$

# Multivariate DLMs

Parameter evolution

$$\boldsymbol{\theta}_t = \mathbf{G}\boldsymbol{\theta}_{t-1} + \mathbf{w}_t \text{ with } \mathbf{w}_t \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

becomes

$$\boldsymbol{\theta}_t = (\mathbf{G} \otimes \mathbf{I}_n) \boldsymbol{\theta}_{t-1} + \mathbf{w}_t \text{ with } \mathbf{w}_t \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

# Multivariate DLMs

## Parameter evolution

If we have 2 regression parameters and  $n = 2$ , then

$$\boldsymbol{\theta}_t = \begin{bmatrix} \alpha_{1,t} \\ \alpha_{2,t} \\ \beta_{1,t} \\ \beta_{2,t} \end{bmatrix} \text{ and } \mathbf{G} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = \mathbf{I}_2$$

# Multivariate DLMs

Parameter evolution

$$\boldsymbol{\theta}_t = (\mathbf{G} \otimes \mathbf{I}_n) \boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$



$$\boldsymbol{\theta}_t = (\mathbf{I}_2 \otimes \mathbf{I}_2) \boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

# Multivariate DLMs

$$\mathbf{I}_m \otimes \mathbf{I}_n = \mathbf{I}_{mn}$$

$$\begin{aligned}\mathbf{I}_2 \otimes \mathbf{I}_2 &= \begin{bmatrix} 1 \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} & 0 \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \\ 0 \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} & 1 \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \end{bmatrix} \\ &= \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}\end{aligned}$$

# Multivariate DLMs

Parameter evolution

$$\boldsymbol{\theta}_t = (\mathbf{G} \otimes \mathbf{I}_n) \boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

$\Downarrow$

$$\begin{bmatrix} \alpha_{1,t} \\ \alpha_{2,t} \\ \beta_{1,t} \\ \beta_{2,t} \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} \alpha_{1,t-1} \\ \alpha_{2,t-1} \\ \beta_{1,t-1} \\ \beta_{2,t-1} \end{bmatrix} + \begin{bmatrix} w_{\alpha_1,t} \\ w_{\alpha_2,t} \\ w_{\beta_1,t} \\ w_{\beta_2,t} \end{bmatrix}$$

$\Downarrow$

$$\boldsymbol{\theta}_t = \boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

# Multivariate DLMs

Evolution variance

$$\boldsymbol{\theta}_t = \boldsymbol{\theta}_{t-1} + \mathbf{w}_t \text{ with } \mathbf{w}_t \sim \text{MVN}(\mathbf{0}, \underline{\mathbf{Q}})$$

What form should we choose for  $\underline{\mathbf{Q}}$ ?

# Multivariate DLMs

Evolution variance

$$\begin{bmatrix} \alpha_t \\ \beta_t \end{bmatrix} \sim \text{MVN} \left( \begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{Q}_\alpha & \mathbf{0} \\ \mathbf{0} & \mathbf{Q}_\beta \end{bmatrix} \right)$$

$$\mathbf{Q}_{(\cdot)} = \begin{bmatrix} q_{(\cdot)} & 0 & \dots & 0 \\ 0 & q_{(\cdot)} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & q_{(\cdot)} \end{bmatrix}$$

Diagonal and equal (IID)

# Multivariate DLMs

Evolution variance

$$\begin{bmatrix} \alpha_t \\ \beta_t \end{bmatrix} \sim \text{MVN} \left( \begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{Q}_\alpha & \mathbf{0} \\ \mathbf{0} & \mathbf{Q}_\beta \end{bmatrix} \right)$$

$$\mathbf{Q}_{(\cdot)} = \begin{bmatrix} q_{(\cdot)1} & 0 & \dots & 0 \\ 0 & q_{(\cdot)2} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & q_{(\cdot)n} \end{bmatrix}$$

Diagonal and unequal

# Multivariate DLMs

Evolution variance

$$\begin{bmatrix} \alpha_t \\ \beta_t \end{bmatrix} \sim \text{MVN} \left( \begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{Q}_\alpha & \mathbf{0} \\ \mathbf{0} & \mathbf{Q}_\beta \end{bmatrix} \right)$$

$$\mathbf{Q}_{(\cdot)} = \begin{bmatrix} q_{(\cdot)1,1} & q_{(\cdot)1,2} & \cdots & q_{(\cdot)1,n} \\ q_{(\cdot)2,1} & q_{(\cdot)2,2} & \cdots & q_{(\cdot)2,n} \\ \vdots & \vdots & \ddots & \vdots \\ q_{(\cdot)n,1} & q_{(\cdot)n,2} & \cdots & q_{(\cdot)n,n} \end{bmatrix}$$

Unconstrained

# Multivariate DLMs

Evolution variance

$$\begin{bmatrix} \alpha_t \\ \beta_t \end{bmatrix} \sim \text{MVN} \left( \begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{Q}_\alpha & \mathbf{0} \\ \mathbf{0} & \mathbf{Q}_\beta \end{bmatrix} \right)$$

In practice, keep  $\mathbf{Q}$  as simple as possible

# Topics for today

## Univariate response

- Stochastic level & growth
- Dynamic Regression
- Dynamic Regression with fixed season
- Forecasting with a DLM
- Model diagnostics

## Multivariate response