

Dynamic Linear Models

FISH 507 – Applied Time Series Analysis

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2 February 2021

Topics for today

Univariate response

- Stochastic level & growth
- Dynamic Regression
- Dynamic Regression with fixed season
- Forecasting with a DLM
- Model diagnostics

Multivariate response

Simple linear regression

Let's begin with a linear regression model

$$y_i = \alpha + \beta x_i + e_i \text{ with } e_i \sim \mathbf{N}(0, \sigma^2)$$

The index i has no explicit meaning in that shuffling (y_i, x_i) pairs has no effect on parameter estimation

Simple linear regression

We can write the model in matrix form

$$y_i = \alpha + \beta x_i + e_i$$
$$\Downarrow$$
$$y_i = \begin{bmatrix} 1 & x_i \end{bmatrix} \begin{bmatrix} \alpha \\ \beta \end{bmatrix} + e_i$$

Simple linear regression

We can write the model in matrix form

$$\begin{aligned}y_i &= \alpha + \beta x_i + e_i \\ &\Downarrow \\ y_i &= \begin{bmatrix} 1 & x_i \end{bmatrix} \begin{bmatrix} \alpha \\ \beta \end{bmatrix} + e_i \\ &\Downarrow \\ y_i &= \mathbf{X}_i^\top \boldsymbol{\theta} + e_i\end{aligned}$$

with

$$\mathbf{X}_i^\top = \begin{bmatrix} 1 & x_i \end{bmatrix} \text{ and } \boldsymbol{\theta} = \begin{bmatrix} \alpha & \beta \end{bmatrix}^\top$$

Dynamic linear model (DLM)

In a *dynamic* linear model, the regression parameters change over time, so we write

$$y_i = \mathbf{X}_i^\top \boldsymbol{\theta} + e_i \quad (\text{static})$$

as

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t \quad (\text{dynamic})$$

Dynamic linear model (DLM)

There are 2 important points here:

$$y_{\boxed{t}} = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t$$

1. Subscript t explicitly acknowledges implicit info in the time ordering of the data in \mathbf{y}

Dynamic linear model (DLM)

There are 2 important points here:

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_{\boxed{t}} + e_t$$

1. Subscript t explicitly acknowledges implicit info in the time ordering of the data in \mathbf{y}
2. The relationship between \mathbf{y} and \mathbf{X} is unique for every t

Constraining a DLM

Close examination of the DLM reveals an apparent problem for parameter estimation

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t$$

Constraining a DLM

Close examination of the DLM reveals an apparent problem for parameter estimation

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t$$

We only have 1 data point per time step (ie, y_t is a scalar)

Thus, we can only estimate 1 parameter (with no uncertainty)!

Constraining a DLM

To address this issue, we'll constrain the regression parameters to be dependent from t to $t + 1$

$$\boldsymbol{\theta}_t = \mathbf{G}_t \boldsymbol{\theta}_{t-1} + \mathbf{w}_t \text{ with } \mathbf{w}_t \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

Constraining a DLM

In practice, we often make \mathbf{G}_t time invariant

$$\boldsymbol{\theta}_t = \mathbf{G}\boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

Constraining a DLM

In practice, we often make \mathbf{G}_t time invariant

$$\boldsymbol{\theta}_t = \mathbf{G}\boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

or assume \mathbf{G}_t is an $m \times m$ identity matrix \mathbf{I}_m

$$\begin{aligned}\boldsymbol{\theta}_t &= \mathbf{I}_m\boldsymbol{\theta}_{t-1} + \mathbf{w}_t \\ &= \boldsymbol{\theta}_{t-1} + \mathbf{w}_t\end{aligned}$$

In the latter case, the parameters follow a random walk over time

DLM in state-space form

Observation model relates the covariates \mathbf{X} to the data

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t$$

State model determines how parameters “evolve” over time

$$\boldsymbol{\theta}_t = \mathbf{G}\boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

DLM in MARSS notation

Full state-space form

$$\begin{aligned}y_t &= \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t \\ \boldsymbol{\theta}_t &= \mathbf{G}\boldsymbol{\theta}_{t-1} + \mathbf{w}_t \\ &\Downarrow \\ y_t &= \mathbf{Z}_t \mathbf{x}_t + v_t \\ \mathbf{x}_t &= \mathbf{B}\mathbf{x}_{t-1} + \mathbf{w}_t\end{aligned}$$

where

$$\mathbf{Z}_t = \mathbf{X}_t^\top, \mathbf{x}_t = \boldsymbol{\theta}_t, v_t = e_t, \mathbf{B} = \mathbf{G}$$

Contrast in covariate effects

Note: DLMs include covariate effect in the observation eqn much differently than other forms of MARSS models

DLM: \mathbf{Z}_t is covariates, \mathbf{x}_t is parameters

$$y_t = \boxed{\mathbf{Z}_t \mathbf{x}_t} + v_t$$

Others: \mathbf{d}_t is covariates, \mathbf{D} is parameters

$$y_t = \mathbf{Z}_t \mathbf{x}_t + \boxed{\mathbf{D} \mathbf{d}_t} + v_t$$

Other forms of DLMs

The regression model is but one type

Others include:

- stochastic “level” (intercept)
- stochastic “growth” (trend, bias)
- seasonal effects (fixed, harmonic)

The most simple DLM

Stochastic level

$$y_t = \alpha_t + e_t$$
$$\alpha_t = \alpha_{t-1} + w_t$$

The most simple DLM

Stochastic level = random walk with obs error

$$y_t = \alpha_t + e_t$$

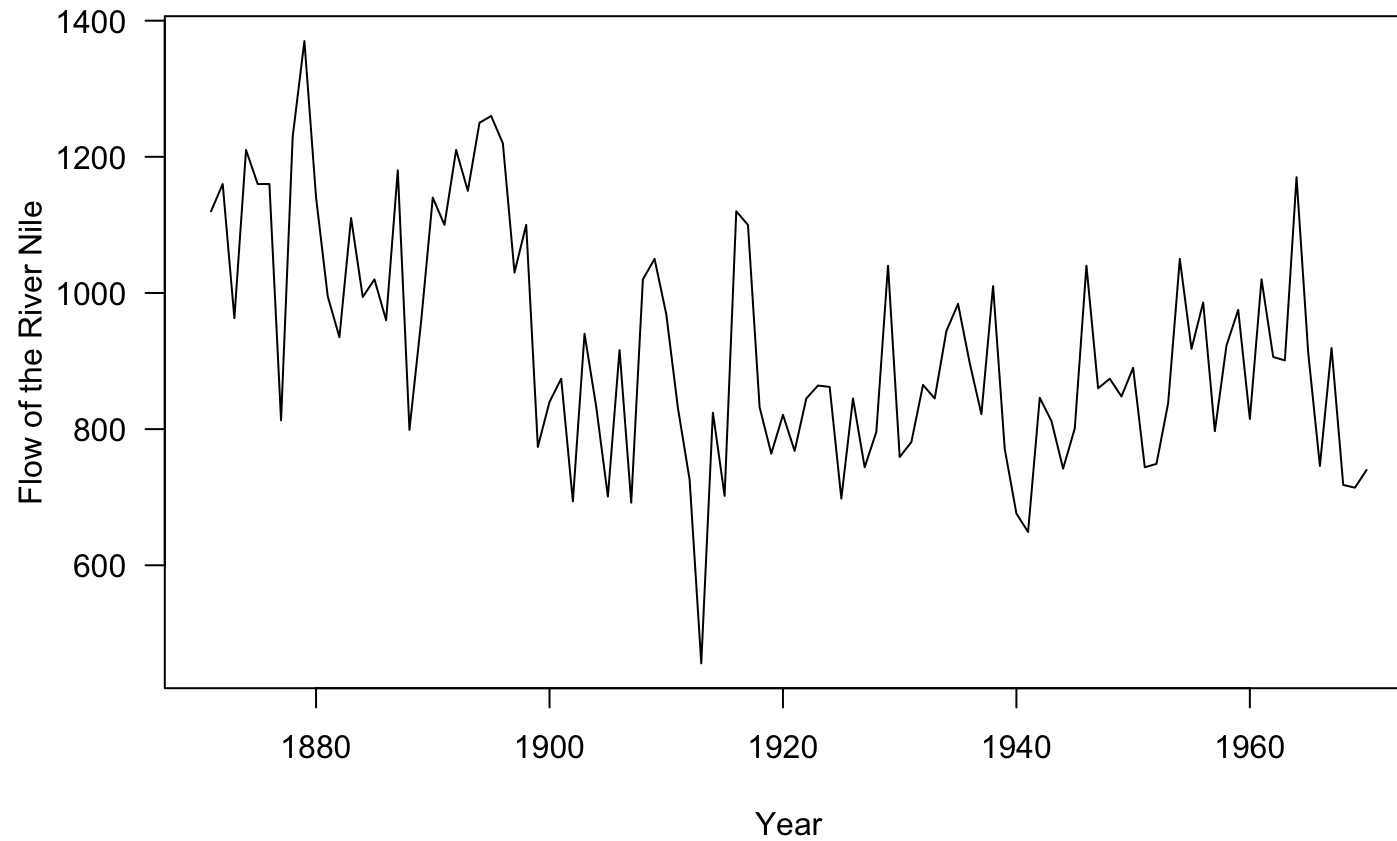
$$\alpha_t = \alpha_{t-1} + w_t$$

⇓

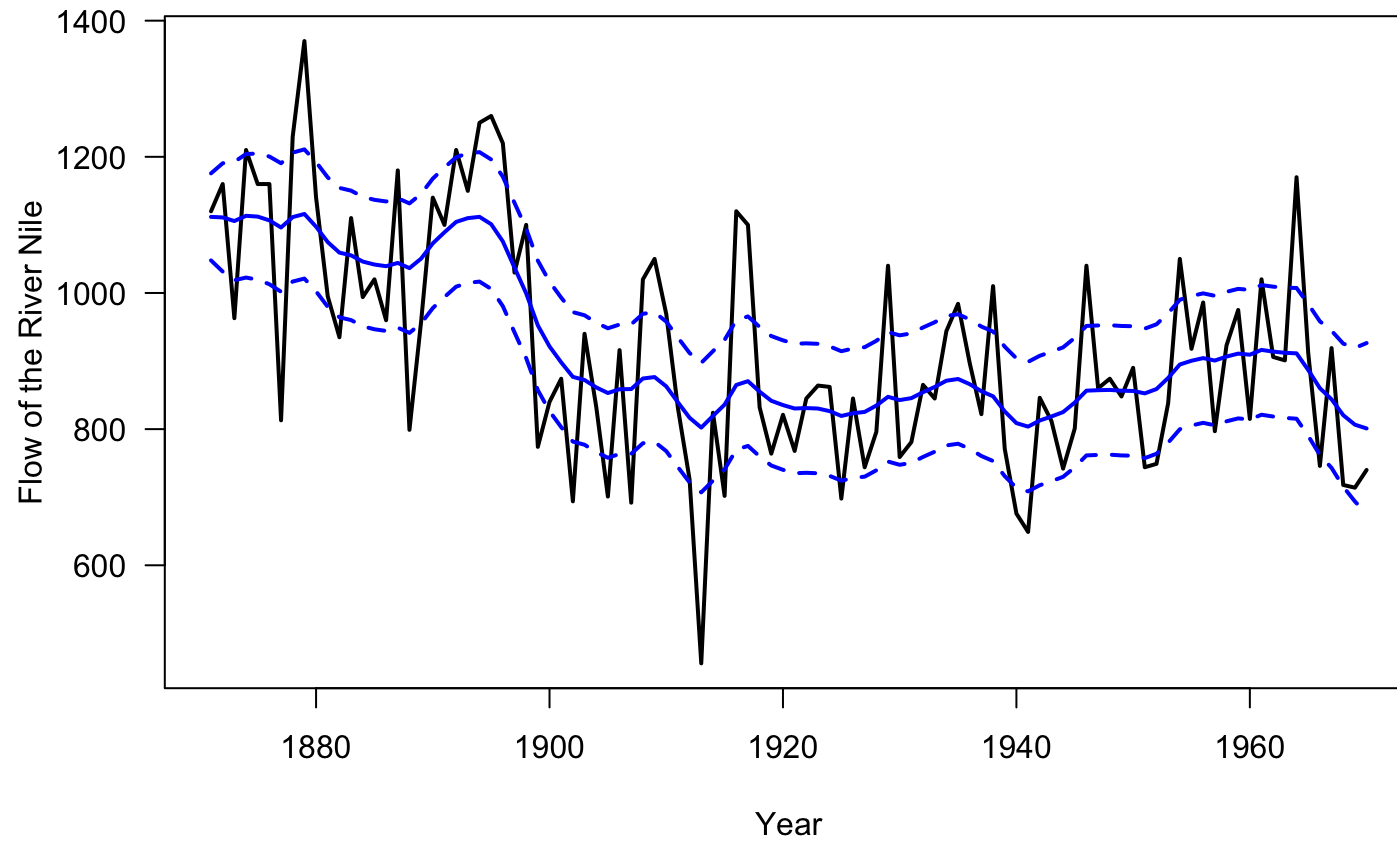
$$y_t = x_t + v_t$$

$$x_t = x_{t-1} + w_t$$

Ex of stochastic level model



Ex of stochastic level model



Univariate DLM for level & growth

Stochastic "level" α_t with deterministic "growth" η

$$y_t = \alpha_t + e_t$$
$$\alpha_t = \alpha_{t-1} + \eta + w_t$$

Univariate DLM for level & growth

Stochastic “level” α_t with deterministic “growth” η

$$\begin{aligned}y_t &= \alpha_t + e_t \\ \alpha_t &= \alpha_{t-1} + \eta + w_t \\ &\Downarrow \\ y_t &= x_t + v_t \\ x_t &= x_{t-1} + u + w_t\end{aligned}$$

This is just a random walk with bias u

Univariate DLM for level & growth

Stochastic “level” α_t with stochastic “growth” η_t

$$y_t = \alpha_t + e_t$$

$$\alpha_t = \alpha_{t-1} + \eta_{t-1} + w_{\alpha,t}$$

$$\eta_t = \eta_{t-1} + w_{\eta,t}$$

Now the “growth” term η_t evolves as well

Univariate DLM for level & growth

Evolution of α_t and η_t

$$\alpha_t = \alpha_{t-1} + \eta_{t-1} + w_{\alpha,t}$$

$$\eta_t = \eta_{t-1} + w_{\eta,t}$$

How do we make this work in practice?

Univariate DLM for level & growth

Evolution of α_t and η_t

$$\alpha_t = \alpha_{t-1} + \eta_{t-1} + w_{\alpha,t}$$

$$\eta_t = \eta_{t-1} + w_{\eta,t}$$

↓

$$\alpha_t = 1\alpha_{t-1} + 1\eta_{t-1} + w_{\alpha,t}$$

$$\eta_t = 0\alpha_{t-1} + 1\eta_{t-1} + w_{\eta,t}$$

Rewrite the equations with explicit coefficients on α_{t-1} and η_{t-1}

Univariate DLM for level & growth

Evolution of α_t and η_t

$$\alpha_t = \alpha_{t-1} + \eta_{t-1} + w_{\alpha,t}$$

$$\eta_t = \eta_{t-1} + w_{\eta,t}$$

\Downarrow

$$\alpha_t = \underline{1}\alpha_{t-1} + \underline{1}\eta_{t-1} + w_{\alpha,t}$$

$$\eta_t = \underline{0}\alpha_{t-1} + \underline{1}\eta_{t-1} + w_{\eta,t}$$

\Downarrow

$$\underbrace{\begin{bmatrix} \alpha_t \\ \eta_t \end{bmatrix}}_{\theta_t} = \underbrace{\begin{bmatrix} \underline{1} & \underline{1} \\ \underline{0} & \underline{1} \end{bmatrix}}_{\mathbf{G}} \underbrace{\begin{bmatrix} \alpha_{t-1} \\ \eta_{t-1} \end{bmatrix}}_{\theta_{t-1}} + \underbrace{\begin{bmatrix} w_{\alpha,t} \\ w_{\eta,t} \end{bmatrix}}_{\mathbf{w}_t}$$

Univariate DLM for level & growth

Evolution of α_t and η_t in MARSS form

$$x_{1,t} = x_{1,t-1} + x_{2,t-1} + w_{1,t}$$

$$x_{2,t} = x_{2,t-1} + w_{2,t}$$

\Downarrow

$$x_{1,t} = \underline{1}x_{1,t-1} + \underline{1}x_{2,t-1} + w_{1,t}$$

$$x_{2,t} = \underline{0}x_{1,t-1} + \underline{1}x_{2,t-1} + w_{2,t}$$

\Downarrow

$$\underbrace{\begin{bmatrix} x_{1,t} \\ x_{2,t} \end{bmatrix}}_{\mathbf{x}_t} = \underbrace{\begin{bmatrix} \underline{1} & \underline{1} \\ \underline{0} & \underline{1} \end{bmatrix}}_{\mathbf{B}} \underbrace{\begin{bmatrix} x_{1,t-1} \\ x_{2,t-1} \end{bmatrix}}_{\mathbf{x}_{t-1}} + \underbrace{\begin{bmatrix} w_{1,t} \\ w_{2,t} \end{bmatrix}}_{\mathbf{w}_t}$$

Univariate DLM for level & growth

Observation model for stochastic *level* & *growth*

$$y_t = \alpha_t + v_t$$

⇓

$$y_t = \underline{1}\alpha_t + \underline{0}\eta_t + v_t$$

Again, rewrite equation with explicit coefficients on α_t and η_t

Univariate DLM for level & growth

Observation model for stochastic *level & growth*

$$y_t = \alpha_t + v_t$$

⇓

$$y_t = \underline{1}\alpha_t + \underline{0}\eta_t + v_t$$

⇓

$$y_t = \underbrace{\begin{bmatrix} \underline{1} & \underline{0} \end{bmatrix}}_{\mathbf{x}_t^\top} \underbrace{\begin{bmatrix} \alpha_t \\ \eta_t \end{bmatrix}}_{\boldsymbol{\theta}_t} + v_t$$

Univariate DLM for level & growth

Obs model for stochastic *level* & *growth* in MARSS form

$$y_t = x_t + v_t$$

⇓

$$y_t = \underline{1}x_{1,t} + \underline{0}x_{2,t} + v_t$$

⇓

$$y_t = \underbrace{\begin{bmatrix} \underline{1} & \underline{0} \end{bmatrix}}_{\mathbf{Z}_t} \underbrace{\begin{bmatrix} x_{1,t} \\ x_{2,t} \end{bmatrix}}_{\mathbf{x}_t} + v_t$$

Univariate DLM for regression

Stochastic intercept and slope

$$y_t = \alpha_t + \beta_t x_t + v_t$$

Univariate DLM for regression

Stochastic intercept and slope

$$y_t = \alpha_t + \beta_t x_t + v_t$$

↓

$$y_t = \underline{1}\alpha_t + \underline{x_t}\beta_t + v_t$$

Rewrite the equation with explicit coefficients for α_t & β_t

Univariate DLM for regression

Stochastic intercept and slope

$$y_t = \alpha_t + \beta_t x_t + v_t$$

⇓

$$y_t = \underline{1}\alpha_t + \underline{x_t}\beta_t + v_t$$

⇓

$$y_t = \underbrace{\begin{bmatrix} \underline{1} & \underline{x_t} \end{bmatrix}}_{\mathbf{x}_t^\top} \underbrace{\begin{bmatrix} \alpha_t \\ \beta_t \end{bmatrix}}_{\boldsymbol{\theta}_t} + v_t$$

Univariate DLM for regression

Stochastic intercept and slope in MARSS form

$$y_t = x_{1,t} + x_{1,t}z_{2,t} + v_t$$

⇓

$$y_t = \underline{1}x_{1,t} + \underline{z_{2,t}}x_{2,t} + v_t$$

⇓

$$y_t = \underbrace{\begin{bmatrix} \underline{1} & \underline{z_{2,t}} \end{bmatrix}}_{\mathbf{Z}_t} \underbrace{\begin{bmatrix} x_{1,t} \\ x_{2,t} \end{bmatrix}}_{\mathbf{x}_t} + v_t$$

Univariate DLM for regression

Parameter evolution follows a random walk

$$\alpha_t = \alpha_{t-1} + w_{\alpha,t}$$

$$\beta_t = \beta_{t-1} + w_{\beta,t}$$

⇓

$$\underbrace{\begin{bmatrix} \alpha_t \\ \beta_t \end{bmatrix}}_{\boldsymbol{\theta}_t} = \underbrace{\begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \end{bmatrix}}_{\boldsymbol{\theta}_{t-1}} + \underbrace{\begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \end{bmatrix}}_{\mathbf{w}_t}$$

Univariate DLM for regression

Parameter evolution in MARSS form

$$x_{1,t} = x_{1,t-1} + w_{1,t}$$

$$x_{2,t} = x_{2,t-1} + w_{2,t}$$

⇓

$$\underbrace{\begin{bmatrix} x_{1,t} \\ x_{2,t} \end{bmatrix}}_{\mathbf{x}_t} = \underbrace{\begin{bmatrix} x_{1,t-1} \\ x_{2,t-1} \end{bmatrix}}_{\mathbf{x}_{t-1}} + \underbrace{\begin{bmatrix} w_{1,t} \\ w_{2,t} \end{bmatrix}}_{\mathbf{w}_t}$$

Univariate DLM with seasonal effect

Dynamic linear regression with fixed seasonal effect

$$y_t = \alpha_t + \beta_t x_t + \gamma_{qtr} + e_t$$

$$\gamma_{qtr} = \begin{cases} \gamma_1 & \text{if } qtr = 1 \\ \gamma_2 & \text{if } qtr = 2 \\ \gamma_3 & \text{if } qtr = 3 \\ \gamma_4 & \text{if } qtr = 4 \end{cases}$$

Univariate DLM with seasonal effect

Dynamic linear regression with fixed seasonal effect

$$y_t = \alpha_t + \beta_t x_t + \gamma_{qtr} + e_t$$

⇓

$$y_t = \begin{bmatrix} 1 & x_t & 1 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_{qtr} \end{bmatrix} + e_t$$

Rewrite the equation with explicit coefficients on parameters

Univariate DLM with seasonal effect

Evolution of parameters

$$\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_{qtr} \end{bmatrix} = \begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ ? \end{bmatrix} + \begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ ? \end{bmatrix}$$

How should we model the fixed effect of γ_{qtr} ?

Univariate DLM with seasonal effect

Evolution of parameters

$$\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_{qtr} \end{bmatrix} = \begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ \gamma_{qtr} \end{bmatrix} + \begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ 0 \end{bmatrix}$$

We don't want γ_{qtr} to evolve as a function of the previous t

Univariate DLM with seasonal effect

Evolution of parameters

$$\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_{qtr} \end{bmatrix} = \begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ \gamma_{qtr} \end{bmatrix} + \begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ 0 \end{bmatrix}$$

OK, so how do we select the right quarterly effect?

Univariate DLM with seasonal effect

Separate out the quarterly effects

$$y_t = \alpha_t + \beta_t x_t + \gamma_{qtr} + e_t$$

↓

$$y_t = \alpha_t + \beta_t x_t + \gamma_1 + \gamma_2 + \gamma_3 + \gamma_4 + e_t$$

Univariate DLM with seasonal effect

Rewrite quarterly effects in matrix notation

$$y_t = \alpha_t + \beta_t x_t + \gamma_1 + \gamma_2 + \gamma_3 + \gamma_4 + e_t$$

⇓

$$y_t = \begin{bmatrix} 1 & x_t & 1 & 1 & 1 & 1 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}$$

But how do we select only the current quarter?

Univariate DLM with seasonal effect

We could set some values in \mathbf{x}_t to 0 ($qtr = 1$)

$$y_t = \begin{bmatrix} 1 & x_t & 1 & 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}$$

⇓

$$y_t = \alpha_t + \beta_t x_t + \gamma_1 + e_t$$

Univariate DLM with seasonal effect

We could set some values in \mathbf{x}_t to 0 ($qtr = 2$)

$$y_t = \begin{bmatrix} 1 & x_t & 0 & 1 & 0 & 0 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}$$

⇓

$$y_t = \alpha_t + \beta_t x_t + \gamma_2 + e_t$$

Univariate DLM with seasonal effect

But *how* would we set the correct 0/1 values?

$$\mathbf{X}_t^\top = [1 \quad x_t \quad ? \quad ? \quad ? \quad ?]$$

Univariate DLM with seasonal effect

We could instead reorder the γ_i within θ_t ($qtr = 1$)

$$y_t = \begin{bmatrix} 1 & x_t & 1 & 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}$$

⇓

$$y_t = \alpha_t + \beta_t x_t + \gamma_1 + e_t$$

Univariate DLM with seasonal effect

We could instead reorder the γ_i within θ_t ($qtr = 2$)

$$y_t = \begin{bmatrix} 1 & x_t & 1 & 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \\ \gamma_1 \end{bmatrix}$$

⇓

$$y_t = \alpha_t + \beta_t x_t + \gamma_2 + e_t$$

Univariate DLM with seasonal effect

But *how* would we shift the γ_i within $\boldsymbol{\theta}_t$?

$$\boldsymbol{\theta}_t = \begin{bmatrix} \alpha_t \\ \beta_t \\ ? \\ ? \\ ? \\ ? \end{bmatrix}$$

Example of non-diagonal \mathbf{G}

We can use a non-diagonal submatrix in the lower right of \mathbf{G} to get the correct quarter effect

$$\mathbf{G} = \left[\begin{array}{cc|cccc} 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ \hline 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 1 & 0 & 0 & 0 \end{array} \right]$$

Evolving parameters

Quarter 1

$$\underbrace{\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}}_{\theta_t} = \underbrace{\begin{bmatrix} 1 & 0 & | & 0 & 0 & 0 & 0 \\ 0 & 1 & | & 0 & 0 & 0 & 0 \\ \hline 0 & 0 & | & 0 & 1 & 0 & 0 \\ 0 & 0 & | & 0 & 0 & 1 & 0 \\ 0 & 0 & | & 0 & 0 & 0 & 1 \\ 0 & 0 & | & 1 & 0 & 0 & 0 \end{bmatrix}}_{\mathbf{G}} \underbrace{\begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ \gamma_4 \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \end{bmatrix}}_{\theta_{t-1}} + \underbrace{\begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}}_{\mathbf{w}_t}$$

Evolving parameters

Quarter 2

$$\underbrace{\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \\ \gamma_1 \end{bmatrix}}_{\theta_t} = \underbrace{\begin{bmatrix} 1 & 0 & | & 0 & 0 & 0 & 0 \\ 0 & 1 & | & 0 & 0 & 0 & 0 \\ \hline 0 & 0 & | & 0 & 1 & 0 & 0 \\ 0 & 0 & | & 0 & 0 & 1 & 0 \\ 0 & 0 & | & 0 & 0 & 0 & 1 \\ 0 & 0 & | & 1 & 0 & 0 & 0 \end{bmatrix}}_{\mathbf{G}} \underbrace{\begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ \gamma_1 \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \end{bmatrix}}_{\theta_{t-1}} + \underbrace{\begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}}_{\mathbf{w}_t}$$

Evolving parameters

Quarter 3

$$\underbrace{\begin{bmatrix} \alpha_t \\ \beta_t \\ \gamma_3 \\ \gamma_4 \\ \gamma_1 \\ \gamma_2 \end{bmatrix}}_{\theta_t} = \underbrace{\begin{bmatrix} 1 & 0 & | & 0 & 0 & 0 & 0 \\ 0 & 1 & | & 0 & 0 & 0 & 0 \\ \hline 0 & 0 & | & 0 & 1 & 0 & 0 \\ 0 & 0 & | & 0 & 0 & 1 & 0 \\ 0 & 0 & | & 0 & 0 & 0 & 1 \\ 0 & 0 & | & 1 & 0 & 0 & 0 \end{bmatrix}}_{\mathbf{G}} \underbrace{\begin{bmatrix} \alpha_{t-1} \\ \beta_{t-1} \\ \gamma_2 \\ \gamma_3 \\ \gamma_4 \\ \gamma_1 \end{bmatrix}}_{\theta_{t-1}} + \underbrace{\begin{bmatrix} w_{\alpha,t} \\ w_{\beta,t} \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}}_{\mathbf{w}_t}$$

Forecasting with a DLM

Forecasting with a DLM

DLMs are often used in a forecasting context where we want a prediction for time t based on the data up through time $t - 1$

Forecasting with a DLM

Pseudo-code

1. get estimate of today's parameters from yesterday's
2. make prediction based on today's parameters & covariates
3. get observation for today
4. update parameters and repeat

Forecasting with a DLM

Step 1: Define the parameters at time $t = 0$

$$\boldsymbol{\theta}_0 | y_0 = \boldsymbol{\pi} + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \boldsymbol{\Lambda})$$

Forecasting with a DLM

Step 1: Define the parameters at time $t = 0$

$$\boldsymbol{\theta}_0|y_0 = \boldsymbol{\pi} + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \boldsymbol{\Lambda})$$

↓

$$E(\boldsymbol{\theta}_0) = \boldsymbol{\pi}$$

Forecasting with a DLM

Step 1: Define the parameters at time $t = 0$

$$\boldsymbol{\theta}_0|y_0 = \boldsymbol{\pi} + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \boldsymbol{\Lambda})$$

↓

$$E(\boldsymbol{\theta}_0) = \boldsymbol{\pi}$$

and

$$\text{Var}(\boldsymbol{\theta}_0) = \text{Var}(\boldsymbol{\pi}) + \text{Var}(\mathbf{w}_1)$$

$$\text{Var}(\boldsymbol{\theta}_0) = \mathbf{0} + \boldsymbol{\Lambda}$$

$$\text{Var}(\boldsymbol{\theta}_0) = \boldsymbol{\Lambda}$$

Forecasting with a DLM

Step 1: Define the parameters at time $t = 0$

$$\boldsymbol{\theta}_0|y_0 = \boldsymbol{\pi} + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \boldsymbol{\Lambda})$$

↓

$$E(\boldsymbol{\theta}_0) = \boldsymbol{\pi}$$

and

$$\text{Var}(\boldsymbol{\theta}_0) = \boldsymbol{\Lambda}$$

↓

$$\boldsymbol{\theta}_0|y_0 \sim \text{MVN}(\boldsymbol{\pi}, \boldsymbol{\Lambda})$$

Forecasting with a DLM

Step 2: Define the parameters at time $t = 1$

$$\boldsymbol{\theta}_1 | y_0 = \mathbf{G}\boldsymbol{\theta}_0 + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

Forecasting with a DLM

Step 2: Define the parameters at time $t = 1$

$$\boldsymbol{\theta}_1 | y_0 = \mathbf{G}\boldsymbol{\theta}_0 + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

↓

$$E(\boldsymbol{\theta}_1) = \mathbf{G}E(\boldsymbol{\theta}_0)$$

$$E(\boldsymbol{\theta}_1) = \mathbf{G}\boldsymbol{\pi}$$

Forecasting with a DLM

Step 2: Define the parameters at time $t = 1$

$$\boldsymbol{\theta}_1 | y_0 = \mathbf{G}\boldsymbol{\theta}_0 + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

↓

$$E(\boldsymbol{\theta}_1) = \mathbf{G}\boldsymbol{\pi}$$

and

$$\text{Var}(\boldsymbol{\theta}_1) = \mathbf{G}\text{Var}(\boldsymbol{\theta}_0)\mathbf{G}^\top + \text{Var}(\mathbf{w}_1)$$

$$\text{Var}(\boldsymbol{\theta}_1) = \mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q}$$

Forecasting with a DLM

Step 2: Define the parameters at time $t = 1$

$$\boldsymbol{\theta}_1 | y_0 = \mathbf{G}\boldsymbol{\theta}_0 + \mathbf{w}_1 \text{ with } \mathbf{w}_1 \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

↓

$$E(\boldsymbol{\theta}_1) = \mathbf{G}\boldsymbol{\pi}$$

and

$$\text{Var}(\boldsymbol{\theta}_1) = \mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q}$$

↓

$$\boldsymbol{\theta}_1 | y_0 \sim \text{MVN}(\mathbf{G}\boldsymbol{\pi}, \mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q})$$

Forecasting with a DLM

Step 3: Make a forecast of y_t at time $t = 1$

$$y_1 | y_0 = \mathbf{X}_1^\top \boldsymbol{\theta}_1 + e_1 \text{ with } e_1 \sim \mathbf{N}(0, R)$$

↓

$$\mathbf{E}(y_1) = \mathbf{X}_1^\top \mathbf{E}(\boldsymbol{\theta}_1)$$

$$\mathbf{E}(y_1) = \mathbf{X}_1^\top \mathbf{G}\boldsymbol{\pi}$$

Forecasting with a DLM

Step 3: Make a forecast of y_t at time $t = 1$

$$y_1 | y_0 = \mathbf{X}_1^\top \boldsymbol{\theta}_1 + e_1 \text{ with } e_1 \sim \mathbf{N}(0, R)$$

↓

$$\mathbf{E}(y_1) = \mathbf{X}_1^\top \mathbf{G}\boldsymbol{\pi}$$

and

$$\text{Var}(y_1) = \mathbf{X}_1^\top \text{Var}(\boldsymbol{\theta}_1) \mathbf{X}_1 + \text{Var}(e_1)$$

$$\text{Var}(y_1) = \mathbf{X}_1^\top [\mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q}] \mathbf{X}_1 + R$$

Forecasting with a DLM

Step 3: Make a forecast of y_t at time $t = 1$

$$y_1 | y_0 = \mathbf{X}_1^\top \boldsymbol{\theta}_1 + e_1 \text{ with } e_1 \sim \mathbf{N}(0, R)$$

↓

$$\mathbf{E}(y_1) = \mathbf{X}_1^\top \mathbf{G}\boldsymbol{\pi}$$

and

$$\text{Var}(y_1) = \mathbf{X}_1^\top [\mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q}]\mathbf{X}_1 + R$$

↓

$$y_1 | y_0 \sim \mathbf{N}(\mathbf{X}_1^\top [\mathbf{G}\boldsymbol{\pi}], \mathbf{X}_1^\top [\mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q}]\mathbf{X}_1 + R)$$

Forecasting with a DLM

Putting it all together

$$\boldsymbol{\theta}_0 | y_0 \sim \text{MVN}(\boldsymbol{\pi}, \boldsymbol{\Lambda})$$

$$\boldsymbol{\theta}_t | y_{t-1} \sim \text{MVN}(\mathbf{G}\boldsymbol{\pi}, \mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q})$$

$$y_t | y_{t-1} \sim \text{N}(\mathbf{X}_t^\top [\mathbf{G}\boldsymbol{\pi}], \mathbf{X}_t^\top [\mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q}]\mathbf{X}_t + R)$$

Forecasting with a DLM

Putting it all together

$$\boldsymbol{\theta}_0 | y_0 \sim \text{MVN}(\boldsymbol{\pi}, \boldsymbol{\Lambda})$$

$$\boldsymbol{\theta}_t | y_{t-1} \sim \text{MVN}(\mathbf{G}\boldsymbol{\pi}, \mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q})$$

$$y_t | y_{t-1} \sim \text{N}(\mathbf{X}_t^\top [\mathbf{G}\boldsymbol{\pi}], \mathbf{X}_t^\top [\mathbf{G}\boldsymbol{\Lambda}\mathbf{G}^\top + \mathbf{Q}]\mathbf{X}_t + R)$$

Using `MARSS()` will make this easy to do

Diagnostics for DLMs

Just as with other models, we'd like to know if our fitted DLM meets its underlying assumptions

We can calculate the forecast error e_t as

$$e_t = y_t - \hat{y}_t$$

and check if

$$(1) e_t \sim \text{N}(0, \sigma)$$

$$(2) \text{Cov}(e_t, e_{t-1}) = 0$$

with a QQ-plot (1) and an ACF (2)

Multivariate DLMS

Multivariate DLMS

We can expand our DLM to have a multivariate response

The most simple multivariate DLM

Multiple observations of a stochastic level

$$\begin{aligned} \mathbf{y}_t &= \mathbf{Z}\alpha_t + \mathbf{v}_t & \mathbf{y}_t \text{ is } n \times 1 \\ \alpha_t &= \alpha_{t-1} + w_t & \alpha_t \text{ is } 1 \times 1 \end{aligned}$$

with

$$\mathbf{Z} = \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix}$$

The most simple multivariate DLM

Multiple observations of a random walk

$$\begin{aligned} \mathbf{y}_t &= \mathbf{Z}x_t + \mathbf{v}_t & \mathbf{y}_t \text{ is } n \times 1 \\ x_t &= x_{t-1} + w_t & x_t \text{ is } 1 \times 1 \end{aligned}$$

with

$$\mathbf{Z} = \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix}$$

Another simple multivariate DLM

Multiple observations of multiple levels

$$\mathbf{y}_t = \mathbf{Z}\boldsymbol{\alpha}_t + \mathbf{v}_t \quad \mathbf{y}_t \text{ is } n \times 1$$

$$\boldsymbol{\alpha}_t = \boldsymbol{\alpha}_{t-1} + \mathbf{w}_t \quad \boldsymbol{\alpha}_t \text{ is } n \times 1$$

with

$$\mathbf{Z} = \mathbf{I}_n = \begin{bmatrix} 1 & 0 & \dots & 0 \\ 0 & 1 & \ddots & 0 \\ \vdots & \ddots & \ddots & 0 \\ 0 & \dots & 0 & 1 \end{bmatrix}$$

Another simple multivariate DLM

Multiple observations of multiple random walks

$$\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{v}_t \quad \mathbf{y}_t \text{ is } n \times 1$$

$$\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t \quad \mathbf{x}_t \text{ is } n \times 1$$

with

$$\mathbf{Z} = \mathbf{I}_n = \begin{bmatrix} 1 & 0 & \dots & 0 \\ 0 & 1 & \ddots & 0 \\ \vdots & \ddots & \ddots & 0 \\ 0 & \dots & 0 & 1 \end{bmatrix}$$

Multivariate DLMs

Regression model

Our univariate model

$$y_t = \mathbf{X}_t^\top \boldsymbol{\theta}_t + e_t \text{ with } e_t \sim \mathbf{N}(0, R)$$

becomes

$$\mathbf{y}_t = (\mathbf{X}_t^\top \otimes \mathbf{I}_n) \boldsymbol{\theta}_t + \mathbf{e}_t \text{ with } \mathbf{e}_t \sim \text{MVN}(\mathbf{0}, \mathbf{R})$$

where \otimes is the *Kronecker product*

Kronecker products

If \mathbf{A} is an $m \times n$ matrix and \mathbf{B} is a $p \times q$ matrix

then $\mathbf{A} \otimes \mathbf{B}$ will be an $mp \times nq$ matrix

$$\mathbf{A} \otimes \mathbf{B} = \begin{bmatrix} a_{11}\mathbf{B} & \dots & a_{1n}\mathbf{B} \\ \vdots & \ddots & \vdots \\ a_{m1}\mathbf{B} & \dots & a_{mn}\mathbf{B} \end{bmatrix}$$

Kronecker products

For example

$$\mathbf{A} = \begin{bmatrix} 1 & 2 \\ 3 & 4 \end{bmatrix} \text{ and } \mathbf{B} = \begin{bmatrix} 2 & 4 \\ 6 & 8 \end{bmatrix}$$

so

$$\mathbf{A} \otimes \mathbf{B} = \begin{bmatrix} 1 \begin{bmatrix} 2 & 4 \\ 6 & 8 \end{bmatrix} & 2 \begin{bmatrix} 2 & 4 \\ 6 & 8 \end{bmatrix} \\ 3 \begin{bmatrix} 2 & 4 \\ 6 & 8 \end{bmatrix} & 4 \begin{bmatrix} 2 & 4 \\ 6 & 8 \end{bmatrix} \end{bmatrix} = \begin{bmatrix} 2 & 4 & 4 & 8 \\ 6 & 8 & 12 & 16 \\ 6 & 12 & 8 & 16 \\ 18 & 24 & 24 & 32 \end{bmatrix}$$

Multivariate DLMs

Regression model with $n = 2$

$$\mathbf{y}_t = (\mathbf{X}_t^\top \otimes \mathbf{I}_n) \boldsymbol{\theta}_t + \mathbf{e}_t$$

\Downarrow

$$\begin{bmatrix} y_{1,t} \\ y_{2,t} \end{bmatrix} = \left(\begin{bmatrix} 1 & x_t \end{bmatrix} \otimes \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \right) \begin{bmatrix} \alpha_{1,t} \\ \alpha_{2,t} \\ \beta_{1,t} \\ \beta_{2,t} \end{bmatrix} + \begin{bmatrix} e_{1,t} \\ e_{2,t} \end{bmatrix}$$

Multivariate DLMs

$$\begin{bmatrix} y_{1,t} \\ y_{2,t} \end{bmatrix} = \left(\begin{bmatrix} 1 & x_t \end{bmatrix} \otimes \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \right) \begin{bmatrix} \alpha_{1,t} \\ \alpha_{2,t} \\ \beta_{1,t} \\ \beta_{2,t} \end{bmatrix} + \begin{bmatrix} e_{1,t} \\ e_{2,t} \end{bmatrix}$$

⇓

$$\begin{bmatrix} y_{1,t} \\ y_{2,t} \end{bmatrix} = \begin{bmatrix} 1 & 0 & x_t & 0 \\ 0 & 1 & 0 & x_t \end{bmatrix} \begin{bmatrix} \alpha_{1,t} \\ \alpha_{2,t} \\ \beta_{1,t} \\ \beta_{2,t} \end{bmatrix} + \begin{bmatrix} e_{1,t} \\ e_{2,t} \end{bmatrix}$$

Multivariate DLMS

Covariance of observation errors

$$\mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma & 0 & 0 & 0 \\ 0 & \sigma & 0 & 0 \\ 0 & 0 & \sigma & 0 \\ 0 & 0 & 0 & \sigma \end{bmatrix} \text{ or } \mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma_1 & 0 & 0 & 0 \\ 0 & \sigma_2 & 0 & 0 \\ 0 & 0 & \sigma_3 & 0 \\ 0 & 0 & 0 & \sigma_4 \end{bmatrix}$$
$$\mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma & \gamma & \gamma & \gamma \\ \gamma & \sigma & \gamma & \gamma \\ \gamma & \gamma & \sigma & \gamma \\ \gamma & \gamma & \gamma & \sigma \end{bmatrix} \text{ or } \mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma_1 & 0 & 0 & 0 \\ 0 & \sigma_2 & 0 & \gamma_{2,4} \\ 0 & 0 & \sigma_3 & 0 \\ 0 & \gamma_{2,4} & 0 & \sigma_4 \end{bmatrix}$$

Multivariate DLMs

Parameter evolution

$$\boldsymbol{\theta}_t = \mathbf{G}\boldsymbol{\theta}_{t-1} + \mathbf{w}_t \text{ with } \mathbf{w}_t \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

becomes

$$\boldsymbol{\theta}_t = (\mathbf{G} \otimes \mathbf{I}_n) \boldsymbol{\theta}_{t-1} + \mathbf{w}_t \text{ with } \mathbf{w}_t \sim \text{MVN}(\mathbf{0}, \mathbf{Q})$$

Multivariate DLMs

Parameter evolution

If we have 2 regression parameters and $n = 2$, then

$$\boldsymbol{\theta}_t = \begin{bmatrix} \alpha_{1,t} \\ \alpha_{2,t} \\ \beta_{1,t} \\ \beta_{2,t} \end{bmatrix} \quad \text{and} \quad \mathbf{G} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = \mathbf{I}_2$$

Multivariate DLMs

Parameter evolution

$$\boldsymbol{\theta}_t = (\mathbf{G} \otimes \mathbf{I}_n) \boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

↓

$$\boldsymbol{\theta}_t = (\mathbf{I}_2 \otimes \mathbf{I}_2) \boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

Multivariate DLMs

$$\mathbf{I}_m \otimes \mathbf{I}_n = \mathbf{I}_{mn}$$

$$\begin{aligned} \mathbf{I}_2 \otimes \mathbf{I}_2 &= \begin{bmatrix} 1 \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} & 0 \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \\ 0 \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} & 1 \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \end{bmatrix} \\ &= \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \end{aligned}$$

Multivariate DLMs

Parameter evolution

$$\boldsymbol{\theta}_t = (\mathbf{G} \otimes \mathbf{I}_n) \boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

\Downarrow

$$\begin{bmatrix} \alpha_{1,t} \\ \alpha_{2,t} \\ \beta_{1,t} \\ \beta_{2,t} \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} \alpha_{1,t-1} \\ \alpha_{2,t-1} \\ \beta_{1,t-1} \\ \beta_{2,t-1} \end{bmatrix} + \begin{bmatrix} w_{\alpha_{1,t}} \\ w_{\alpha_{2,t}} \\ w_{\beta_{1,t}} \\ w_{\beta_{2,t}} \end{bmatrix}$$

\Downarrow

$$\boldsymbol{\theta}_t = \boldsymbol{\theta}_{t-1} + \mathbf{w}_t$$

Multivariate DLMS

Evolution variance

$$\boldsymbol{\theta}_t = \boldsymbol{\theta}_{t-1} + \mathbf{w}_t \text{ with } \mathbf{w}_t \sim \text{MVN}(\mathbf{0}, \underline{\mathbf{Q}})$$

What form should we choose for \mathbf{Q} ?

Multivariate DLMs

Evolution variance

$$\begin{bmatrix} \boldsymbol{\alpha}_t \\ \boldsymbol{\beta}_t \end{bmatrix} \sim \text{MVN} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{Q}_\alpha & \mathbf{0} \\ \mathbf{0} & \mathbf{Q}_\beta \end{bmatrix} \right)$$

$$\mathbf{Q}_{(\cdot)} = \begin{bmatrix} q_{(\cdot)} & 0 & \dots & 0 \\ 0 & q_{(\cdot)} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & q_{(\cdot)} \end{bmatrix}$$

Diagonal and equal (IID)

Multivariate DLMs

Evolution variance

$$\begin{bmatrix} \boldsymbol{\alpha}_t \\ \boldsymbol{\beta}_t \end{bmatrix} \sim \text{MVN} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{Q}_\alpha & \mathbf{0} \\ \mathbf{0} & \mathbf{Q}_\beta \end{bmatrix} \right)$$

$$\mathbf{Q}_{(\cdot)} = \begin{bmatrix} q_{(\cdot)1} & 0 & \dots & 0 \\ 0 & q_{(\cdot)2} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & q_{(\cdot)n} \end{bmatrix}$$

Diagonal and unequal

Multivariate DLMs

Evolution variance

$$\begin{bmatrix} \boldsymbol{\alpha}_t \\ \boldsymbol{\beta}_t \end{bmatrix} \sim \text{MVN} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{Q}_\alpha & \mathbf{0} \\ \mathbf{0} & \mathbf{Q}_\beta \end{bmatrix} \right)$$

$$\mathbf{Q}_{(\cdot)} = \begin{bmatrix} q_{(\cdot)1,1} & q_{(\cdot)1,2} & \cdots & q_{(\cdot)1,n} \\ q_{(\cdot)2,1} & q_{(\cdot)2,2} & \cdots & q_{(\cdot)2,n} \\ \vdots & \vdots & \ddots & \vdots \\ q_{(\cdot)n,1} & q_{(\cdot)n,2} & \cdots & q_{(\cdot)n,n} \end{bmatrix}$$

Unconstrained

Multivariate DLMs

Evolution variance

$$\begin{bmatrix} \boldsymbol{\alpha}_t \\ \boldsymbol{\beta}_t \end{bmatrix} \sim \text{MVN} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{Q}_\alpha & \mathbf{0} \\ \mathbf{0} & \mathbf{Q}_\beta \end{bmatrix} \right)$$

In practice, keep \mathbf{Q} as simple as possible

Topics for today

Univariate response

- Stochastic level & growth
- Dynamic Regression
- Dynamic Regression with fixed season
- Forecasting with a DLM
- Model diagnostics

Multivariate response