# **Dynamic Factor Analysis**

FISH 507 – Applied Time Series Analysis

Mark Scheuerell 31 Jan 2019

## **Topics for today**

Deterministic vs stochastic elements

Regression with autocorrelated errors

Regression with temporal random effects

Dynamic Factor Analysis (DFA)

- Forms of covariance matrix
- Constraints for model fitting
- Interpretation of results

Consider this simple model, consisting of a mean  $\mu$  plus error

 $y_i = \mu + e_i$  with  $e_i \sim N(0, \sigma^2)$ 

The right-hand side of the equation is composed of *deterministic* and *stochastic* pieces

$$y_i = \underbrace{\mu}_{\text{deterministic}} + \underbrace{e_i}_{\text{stochastic}}$$

Sometime these pieces are referred to as *fixed* and *random* 

 $y_i = \underbrace{\mu}_{\text{fixed}} + \underbrace{e_i}_{\text{random}}$ 

This can also be seen by rewriting the model

$$y_i = \mu + e_i$$
 with  $e_i \sim N(0, \sigma^2)$ 

as

 $y_i \sim N(\mu, \sigma^2)$ 

## Simple linear regression

We can expand the deterministic part of the model, as with linear regression

$$y_i = \underbrace{\alpha + \beta x_i}_{\text{mean}} + e_i \text{ with } e_i \sim N(0, \sigma^2)$$

mean

SO

$$y_i \sim N(\alpha + \beta x_i, \sigma^2)$$

## A simple time series model

Consider a simple model with a mean  $\mu$  plus white noise

 $y_t = \mu + e_t$  with  $e_t \sim N(0, \sigma^2)$ 

#### Time series model with covariates

We can expand the deterministic part of the model, as before with linear regression

$$y_t = \underbrace{\alpha + \beta x_t}_{\text{mean}} + e_t \text{ with } e_t \sim N(0, \sigma^2)$$

SO

$$y_t \sim N(\alpha + \beta x_t, \sigma^2)$$

#### **Example of linear model**



Time

#### **Model residuals**



These do *not* look like white noise!

#### ACF of model residuals



There is significant autocorrelation at lag = 1

## Model with autocorrelated errors

We can expand the stochastic part of the model to have autocorrelated errors

 $y_t = \alpha + \beta x_t + e_t$  $e_t = \phi e_{t-1} + w_t$ 

with  $w_t \sim N(0, \sigma^2)$ 

## Model with autocorrelated errors

We can expand the stochastic part of the model to have autocorrelated errors

$$y_t = \alpha + \beta x_t + e_t$$
$$e_t = \phi e_{t-1} + w_t$$

with  $w_t \sim N(0, \sigma^2)$ 

We can write this model as our standard state-space model

Observation equation

$$y_t = \alpha + \beta x_t + e_t$$
  
=  $e_t + \alpha + \beta x_t$   
 $\Downarrow$   
 $y_t = x_t + a + Dd_t + v_t$ 

with

$$x_t = e_t, a = \alpha, D = \beta, d_t = x_t, v_t = 0$$

State equation

$$e_t = \phi e_{t-1} + w_t$$
$$\Downarrow$$
$$x_t = Bx_t + w_t$$

with

 $x_t = e_t$  and  $B = \phi$ 

Full form

$$y_t = \alpha + \beta x_t + e_t$$
$$e_t = \phi e_{t-1} + w_t$$
$$\Downarrow$$
$$y_t = a + Dd_t + x_t$$
$$x_t = Bx_t + w_t$$

Observation model in MARSS()

$$y_t = a + Dd_t + x_t$$

$$\Downarrow$$

$$y_t = Zx_t + a + Dd_t + v_t$$

У	= data	## [1 x T] matrix of data
а	<pre>= matrix("a")</pre>	## intercept
D	<pre>= matrix("D")</pre>	## slope
d	= covariate	## [1 x T] matrix of measured covariate
Z	= matrix(1)	## no multiplier on x
R	= matrix(0)	## $v_t \sim N(0,R)$ ; want $y_t = 0$ for all t

State model in MARSS()

$$x_t = Bx_t + w_t$$
$$\Downarrow$$
$$x_t = Bx_t + u + Cc_t + w_t$$

В	=	<pre>matrix("b")</pre>	## AR(1)	coefficient	for	model	errors
Q	=	<pre>matrix("q")</pre>	## w_t ~	N(0,Q); var	for	model	errors
u	=	<pre>matrix(0)</pre>	## u = 0				
С	=	<pre>matrix(0)</pre>	## C = 0				
С	=	<pre>matrix(0)</pre>	## c_t =	0 for all t			

## MORE RANDOM EFFECTS

## Expanding the random effect

Recall our simple model

$$y_t = \underbrace{\mu}_{\text{fixed}} + \underbrace{e_t}_{\text{random}}$$

## Expanding the random effect

We can expand the random portion

$$y_t = \underbrace{\mu}_{\text{fixed}} + \underbrace{f_t + e_t}_{\text{random}}$$

$$e_t \sim N(0, \sigma)$$
  
 $f_t \sim N(f_{t-1}, \gamma)$ 

## Expanding the random effect

We can expand the random portion

$$y_t = \underbrace{\mu}_{\text{fixed}} + \underbrace{f_t + e_t}_{\text{random}}$$

 $e_t \sim N(0, \sigma)$  $f_t \sim N(f_{t-1}, \gamma)$ 

This is simply a random walk observed with error

#### Random walk observed with error

$$y_t = \mu + f_t + e_t \text{ with } e_t \sim N(0, \sigma)$$
  

$$f_t = f_{t-1} + w_t \text{ with } w_t \sim N(0, \gamma)$$
  

$$\Downarrow$$
  

$$y_t = a + x_t + v_t \text{ with } v_t \sim N(0, R)$$
  

$$x_t = x_{t-1} + w_t \text{ with } w_t \sim N(0, Q)$$

## **Expanding fixed & random effects**

We can expand the fixed portion

 $y_t = \underbrace{\alpha + \beta x_t}_{\text{fixed}} + \underbrace{f_t + e_t}_{\text{random}}$  $e_t \sim N(0, \sigma)$  $f_t \sim N(f_{t-1}, \gamma)$ 

#### Fixed & random effects

In familiar state-space form

$$y_t = \alpha + \beta x_t + f_t + e_t \text{ with } e_t \sim N(0, \sigma)$$
  

$$f_t = f_{t-1} + w_t \text{ with } w_t \sim N(0, \gamma)$$
  

$$\Downarrow$$
  

$$y_t = a + Dd_t + x_t + v_t \text{ with } v_t \sim N(0, R)$$
  

$$x_t = x_{t-1} + w_t \text{ with } w_t \sim N(0, Q)$$

## MULTIPLE TIME SERIES

## Simple model for 2+ time series

Random walk observed with error

 $y_{i,t} = x_{i,t} + a_i + v_{i,t}$  $x_{i,t} = x_{i,t-1} + w_{i,t}$ 

with

 $v_{i,t} \sim \mathbf{N}(0, R)$ 

 $w_{i,t} \sim \mathcal{N}(0, Q)$ 

#### Random walk observed with error

$$y_{1,t} = x_{1,t} + a_1 + v_{1,t}$$
  

$$y_{2,t} = x_{2,t} + a_2 + v_{2,t}$$
  

$$\vdots$$
  

$$y_{n,t} = x_{n,t} + a_2 + v_{n,t}$$

$$x_{1,t} = x_{1,t-1} + w_{1,t}$$
  

$$x_{2,t} = x_{2,t-1} + w_{2,t}$$
  

$$\vdots$$
  

$$x_{n,t} = x_{n,t-1} + w_{n,t}$$

# Random walk observed with error

In matrix form

$$\mathbf{y}_t = \mathbf{x}_t + \mathbf{a} + \mathbf{v}_t$$
$$\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t$$

with

 $\mathbf{v}_t \sim \text{MVN}(\mathbf{0}, \mathbf{R})$ 

 $\mathbf{w}_t \sim \mathrm{MVN}(\mathbf{0}, \mathbf{Q})$ 

m mm hall my m My man man when when man I M M M M M M V MA INA MA MA m ~ mm han mm

## **Environmental time series**

We often observe covariance among environmental time series, especially for those close to one another

mm l  $\gamma$  $\mathbb{N}$ M h /M

Are there some common patterns here?

#### Common patterns in time series



Ex: population structure

$$\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{v}_t$$
$$\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t$$

Ex: Harbor seal population structure

$$\begin{bmatrix} y_{1} \\ y_{2} \\ y_{3} \\ y_{4} \\ y_{5} \end{bmatrix}_{t} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \end{bmatrix} \times \begin{bmatrix} x_{JF} \\ x_{N} \\ x_{S} \end{bmatrix}_{t} + \begin{bmatrix} a_{1} \\ a_{2} \\ a_{3} \\ a_{4} \\ a_{5} \end{bmatrix} + \begin{bmatrix} v_{1} \\ v_{2} \\ v_{3} \\ v_{4} \\ v_{5} \end{bmatrix}_{t}$$
$$\begin{bmatrix} x_{JF} \\ x_{N} \\ x_{S} \end{bmatrix}_{t} = \begin{bmatrix} x_{JF} \\ x_{N} \\ x_{S} \end{bmatrix}_{t-1} + \begin{bmatrix} w_{JF} \\ w_{N} \\ w_{S} \end{bmatrix}_{t}$$
# Finding common patterns

What if our observations were instead a mixture of 2+ states?

For example, we sampled haul-outs located between several breeding sites

#### **Mixtures of states**

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \\ y_5 \end{bmatrix}_t = \begin{bmatrix} 0.8 & 0.2 & 0 \\ 0.2 & 0.7 & 0.1 \\ 0 & 0.9 & 0.1 \\ 0 & 0.3 & 0.7 \\ 0 & 0.1 & 0.9 \end{bmatrix} \times \begin{bmatrix} x_{JF} \\ x_N \\ x_S \end{bmatrix}_t + \begin{bmatrix} a_1 \\ a_2 \\ a_3 \\ a_4 \\ a_5 \end{bmatrix} + \begin{bmatrix} v_1 \\ v_2 \\ v_3 \\ v_4 \\ v_5 \end{bmatrix}_t$$
$$\begin{bmatrix} x_{JF} \\ x_N \\ x_S \end{bmatrix}_t = \begin{bmatrix} x_{JF} \\ x_N \\ x_S \end{bmatrix}_t + \begin{bmatrix} w_{JF} \\ w_N \\ w_S \end{bmatrix}_t$$

# Finding common patterns

What if our observations were a mixture of states, but we didn't know how many or the weightings?

$$\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{v}_t$$
$$\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t$$

What are the dimensions of  $\mathbb{Z}$ ?

What are the elements within  $\mathbb{Z}$ ?

DFA is a *dimension reduction* technique, which models *n* observed time series as a function of *m* hidden states (patterns), where  $n \gg m$ 

#### State-space form

$$\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{v}_t$$
$$\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t$$

data:  $\mathbf{y}_t$  is  $n \times 1$ 

loadings: **Z** is  $n \times m$  with n > m

states:  $\mathbf{x}_t$  is  $m \times 1$ 

## **Dimension reduction**

Principal Components Analysis (PCA)

Goal is to reduce some large number of correlated variates into a few uncorrelated factors

Calculating the principal components requires us to estimate the covariance of the data

PC = eigenvectors(cov(y))

There will be *n* principal components (eigenvectors) for an  $n \times T$  matrix **y** 

We reduce the dimension by selecting a subset of the components that explain much of the variance (eg, the first 2)







PC1

#### Relationship between PCA & DFA

We need to estimate the covariance in the data  $\boldsymbol{y}$ 

 $\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{v}_t$ , with  $\mathbf{v}_t \sim \text{MVN}(\mathbf{0}, \mathbf{R})$ 

SO

$$\operatorname{cov}(\mathbf{y}_t) = \mathbf{Z}\operatorname{cov}(\mathbf{x}_t)\mathbf{Z}^\top + \mathbf{R}$$

#### In PCA, we require **R** to be diagonal, but not so in DFA

#### **Principal Components Analysis (PCA)** Forms for **R** with n = 4

$$\mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma & 0 & 0 & 0 \\ 0 & \sigma & 0 & 0 \\ 0 & 0 & \sigma & 0 \\ 0 & 0 & 0 & \sigma \end{bmatrix} \text{ or } \mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma_1 & 0 & 0 & 0 \\ 0 & \sigma_2 & 0 & 0 \\ 0 & 0 & \sigma_3 & 0 \\ 0 & 0 & 0 & \sigma_4 \end{bmatrix}$$

Forms for **R** with n = 4

$$\mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma & 0 & 0 & 0 \\ 0 & \sigma & 0 & 0 \\ 0 & 0 & \sigma & 0 \\ 0 & 0 & \sigma & \sigma \end{bmatrix} \text{ or } \mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma_{1} & 0 & 0 & 0 \\ 0 & \sigma_{2} & 0 & 0 \\ 0 & 0 & \sigma_{3} & 0 \\ 0 & 0 & 0 & \sigma_{4} \end{bmatrix}$$
$$\mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma & \gamma & \gamma & \gamma \\ \gamma & \sigma & \gamma & \gamma \\ \gamma & \gamma & \sigma & \gamma \\ \gamma & \gamma & \sigma & \gamma \\ \gamma & \gamma & \gamma & \sigma \end{bmatrix} \text{ or } \mathbf{R} \stackrel{?}{=} \begin{bmatrix} \sigma_{1} & 0 & 0 & 0 \\ 0 & \sigma_{2} & 0 & \gamma_{2,4} \\ 0 & 0 & \sigma_{3} & 0 \\ 0 & \gamma_{2,4} & 0 & \sigma_{4} \end{bmatrix}$$

$$\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{v}_t$$
$$\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t$$

What form should we use for **Z**?

$$\mathbf{Z} \stackrel{?}{=} \begin{bmatrix} z_1 \\ z_2 \\ z_3 \\ z_4 \\ z_5 \end{bmatrix} \text{ or } \mathbf{Z} \stackrel{?}{=} \begin{bmatrix} z_{1,1} & z_{2,1} \\ z_{1,2} & z_{2,2} \\ z_{1,3} & z_{2,3} \\ z_{1,4} & z_{2,4} \\ z_{1,5} & z_{2,5} \end{bmatrix} \text{ or } \mathbf{Z} \stackrel{?}{=} \begin{bmatrix} z_{1,1} & z_{2,1} & z_{3,1} \\ z_{1,2} & z_{2,2} & z_{3,2} \\ z_{1,3} & z_{2,3} & z_{3,3} \\ z_{1,4} & z_{2,4} \\ z_{1,5} & z_{2,5} \end{bmatrix}$$

$$\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{v}_t$$
$$\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t$$

What form should we use for **Z**?

$$\mathbf{Z} \stackrel{?}{=} \begin{bmatrix} z_1 \\ z_2 \\ z_3 \\ \vdots \\ z_5 \end{bmatrix} \text{ or } \mathbf{Z} \stackrel{?}{=} \begin{bmatrix} z_{1,1} & z_{2,1} \\ z_{1,2} & z_{2,2} \\ z_{1,3} & z_{2,3} \\ \vdots & \vdots \\ z_{1,n} & z_{2,n} \end{bmatrix} \text{ or } \mathbf{Z} \stackrel{?}{=} \begin{bmatrix} z_{1,1} & z_{2,1} & z_{3,1} \\ z_{1,2} & z_{2,2} & z_{3,2} \\ z_{1,3} & z_{2,3} & z_{3,3} \\ \vdots & \vdots \\ z_{1,n} & z_{2,n} \end{bmatrix}$$

We'll use model selection criteria to choose (eg, AICc)

# Fitting DFA models

It turns out that there are an infinite number of combinations of  ${\bf Z}$  and  ${\bf x}$  that will equal  ${\bf y}$ 

Therefore we need to impose some constraints on the model

#### The offset **a**

$$\mathbf{y}_{t} = \mathbf{Z}\mathbf{x}_{t} + \mathbf{a} + \mathbf{v}_{t}$$
$$\mathbf{x}_{t} = \mathbf{x}_{t-1} + \mathbf{w}_{t}$$
$$\mathbf{a} = \begin{bmatrix} a_{1} \\ a_{2} \\ a_{3} \\ \vdots \\ a_{n} \end{bmatrix}$$

We will set the first *m* elements of **a** to 0

The offset **a** 

For example, if n = 5 and m = 2

	$\begin{bmatrix} a_1 \end{bmatrix}$		$\begin{bmatrix} 0 \end{bmatrix}$
-	$a_2$		0
<b>a</b> =	<i>a</i> <sub>3</sub>	$\Rightarrow$ a =	<i>a</i> 3
	<i>a</i> 4		<i>a</i> 4
	$\lfloor a_5 \rfloor$		$\lfloor a_5 \rfloor$

The offset  $\boldsymbol{a}$ 

For example, if n = 5 and m = 2

$$\mathbf{a} = \begin{bmatrix} a_1 \\ a_2 \\ a_3 \\ a_4 \\ a_5 \end{bmatrix} \Rightarrow \mathbf{a} = \begin{bmatrix} 0 \\ 0 \\ a_3 \\ a_4 \\ a_5 \end{bmatrix} \Rightarrow \mathbf{a} = \begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

Note, however, that this causes problems for the EM algorithm so we will often de-mean the data and set  $a_i = 0$  for all *i* 

The loadings  ${\bf Z}$ 

$$\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{v}_t$$
$$\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t$$
$$\begin{bmatrix} z_{1,1} & z_{2,1} & \dots & z_n \end{bmatrix}$$

$$\mathbf{Z} = \begin{bmatrix} z_{1,1} & z_{2,1} & \dots & z_{m,1} \\ z_{1,2} & z_{2,2} & \dots & z_{m,2} \\ z_{1,3} & z_{2,3} & \dots & z_{m,3} \\ \vdots & \vdots & \ddots & z_{m,4} \\ z_{1,n} & z_{2,n} & \dots & z_{m,n} \end{bmatrix}$$

We will set the upper right triangle of  $\mathbf{Z}$  to 0

The loadings  ${\bf Z}$ 

For example, if n = 5 and m = 3

$$\mathbf{Z} = \begin{bmatrix} z_{1,1} & 0 & 0 \\ z_{1,2} & z_{2,2} & 0 \\ z_{1,3} & z_{2,3} & z_{3,3} \\ z_{1,4} & z_{2,3} & z_{3,4} \\ z_{1,5} & z_{2,5} & z_{3,5} \end{bmatrix}$$

For the first m - 1 rows of **Z**,  $z_{i,j} = 0$  if j > i

The loadings  ${\bf Z}$ 

An additional constraint is necessary in a Bayesian context

	Z1,1	0	0
	<i>Z</i> 1,2	<i>Z</i> 2,2	0
<b>Z</b> =	Z1,3	Z2,3	<i>Z</i> 3,3
	<i>Z</i> 1,4	<i>Z</i> 2,3	<i>Z</i> 3,4
	$_{21,5}$	<i>Z</i> 2,5	<i>Z</i> 3,5 _

Diagonal of **Z** is positive:  $z_{i,j} > 0$  if i = j

#### The state variance **Q**

$$\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{v}_t$$
$$\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t$$

 $\mathbf{w}_t \sim \mathrm{MVN}(\mathbf{0}, \mathbf{Q})$ 

We will set  ${f Q}$  equal to the Identity matrix  ${f I}$ 

The state variance  ${f Q}$ 

For example, if m = 4

$$\mathbf{Q} = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

This allows our random walks to have a *lot* of flexibility

Including *p* covariates

 $\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{D}\mathbf{d}_t + \mathbf{v}_t$  $\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t$ 

 $\mathbf{d}_t$  is a  $p \times 1$  vector of covariates at time t

**D** is an  $n \times p$  matrix of covariate effects

#### **Dynamic Factor Analysis (DFA)** Form for **D**

 $\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{D}\mathbf{d}_t + \mathbf{v}_t$  $\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t$ 

Careful thought must be given *a priori* as to the form for **D** 

Should the effect(s) vary by site, species, etc?

## **Dynamic Factor Analysis (DFA)** Form for **D**

For example, given 2 covariates, Temp and Salinity



## A note on model selection

Earlier we saw that we could use model selection criteria to help us choose among the different forms for  ${\bf Z}$ 

However, caution must be given when comparing models with and without covariates, and varying numbers of states

## A note on model selection

Think about the model form

$$\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{D}\mathbf{d}_t + \mathbf{v}_t$$

 $\mathbf{x}_t$  is an *undetermined* random walk

**d**<sub>*t*</sub> is a *predetermined* covariate

Unless  $\mathbf{d}$  is highly correlated with  $\mathbf{y}$ , then the inclusion of a state  $\mathbf{x}$  will be favored over  $\mathbf{d}$ 

## A note on model selection

Thus, work out fixed effects (covariates) while keeping the random effects (states) constant, and vice versa

For example, compare data support for models with different combinations of covariates, only one state (m = 1), and a "diagonal and equal" **R** 

## **Interpreting DFA results**

Recall that we had to constrain the form of  ${\bf Z}$  to fit the model

Z =	Z1,1	0	• • •	0
	<i>Z</i> 1,2	<i>Z</i> 2,2	•.	0
	• •	•	•••	0
	• •	• •	• •	Zm,m
	• •	• •	• •	• •
	$_{21,n}$	$Z_{2,n}$	Z3,n	Zm,n

So, the 1st common factor is determined by the 1st variate, the 2nd common factor by the first two variates, etc.

## **Interpreting DFA results**

To help with this, we can use a *basis rotation* to maximize the loadings on a few factors

If **H** is an  $m \times m$  non-singular matrix, these 2 DFA models are equivalent

 $\mathbf{y}_t = \mathbf{Z}\mathbf{x}_t + \mathbf{a} + \mathbf{D}\mathbf{d}_t + \mathbf{v}_t$  $\mathbf{x}_t = \mathbf{x}_{t-1} + \mathbf{w}_t$ 

$$\mathbf{y}_t = \mathbf{Z}\mathbf{H}^{-1}\mathbf{x}_t + \mathbf{a} + \mathbf{D}\mathbf{d}_t + \mathbf{v}_t$$
$$\mathbf{H}\mathbf{x}_t = \mathbf{H}\mathbf{x}_{t-1} + \mathbf{H}\mathbf{w}_t$$

How should we choose **H**?

#### **Basis rotation**

Varimax

A varimax rotation will maximize the variance of the loadings in  ${\bf Z}$  along a few of the factors

#### PCA of 5 wines with 8 attributes



### **Rotated loadings**



## **Rotated loadings**


## **Topics for today**

Deterministic vs stochastic elements

Regression with autocorrelated errors

Regression with temporal random effects

Dynamic Factor Analysis (DFA)

- Forms of covariance matrix
- Constraints for model fitting
- Interpretation of results