

# Bayesian estimation for time series models

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# Re-cap following models using Bayesian code

- Regression
- ARMA models
- State Space Models
- Dynamic Factor Analysis
- Dynamic Linear Models
- MARSS models (multivariate time series models)

# Why Bayesian?

- Complex hierarchical models
- Inference: what's the probability that the data are less than some threshold?
- No bootstrapping!
  - We get credible intervals for parameters and states simultaneously

# Getting started

```
library(rstan)
```

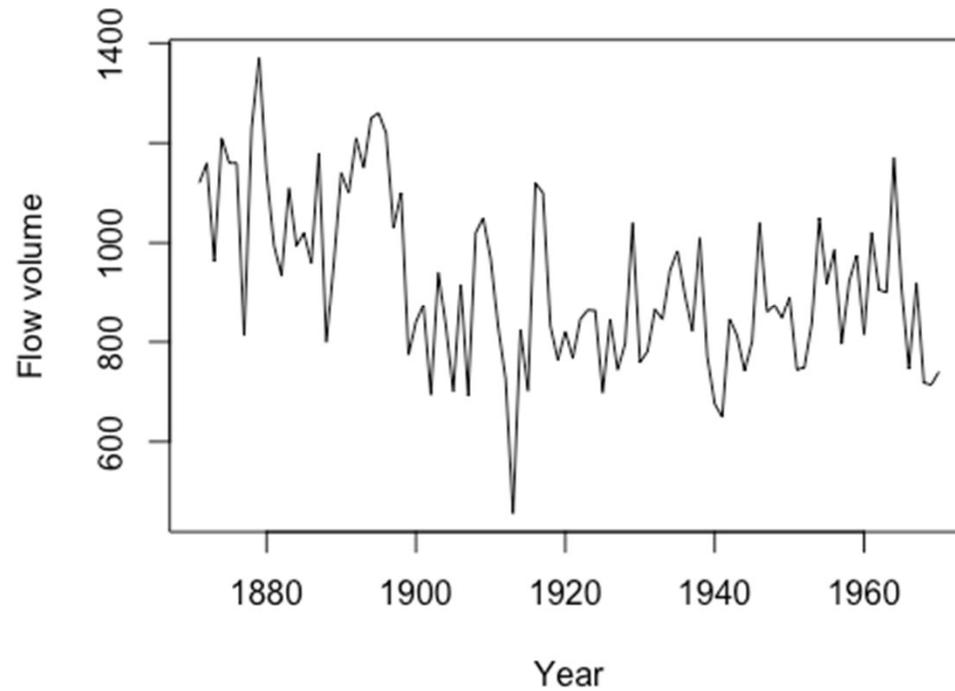
```
library(devtools)
```

```
devtools::install_github("eric-ward/safs-  
timeseries/statss")
```

```
library(statss)
```

# Using STAN objects

- Flow volume from Nile River



# Fitting linear regression

- We'll use wrapper function **'fit\_stan'**

```
x = model.matrix(lm(Nile~1))
```

```
lm_intercept = fit_stan(y = as.numeric(Nile), x =  
rep(1, length(Nile)), model_name = "regression")
```

# Output of fitted object

```
> lm_intercept
```

```
Inference for Stan model: regression.
```

```
3 chains, each with iter=1000; warmup=500; thin=1;
```

```
post-warmup draws per chain=500, total post-warmup draws=1500.
```

	mean	se_mean	sd	2.5%	25%	50%	75%	97.5%	n_eff	Rhat
beta[1]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
sigma	936.34	1.79	66.16	819.09	891.53	929.75	978.62	1074.10	1365	1
pred[1]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
pred[2]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
pred[3]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
pred[4]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
pred[5]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
pred[6]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
pred[7]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
pred[8]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
pred[9]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
pred[10]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
pred[11]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1
pred[12]	0.56	0.05	1.97	-3.25	-0.78	0.51	1.87	4.53	1321	1

# Output from model

- Can be plotted using base graphics in R

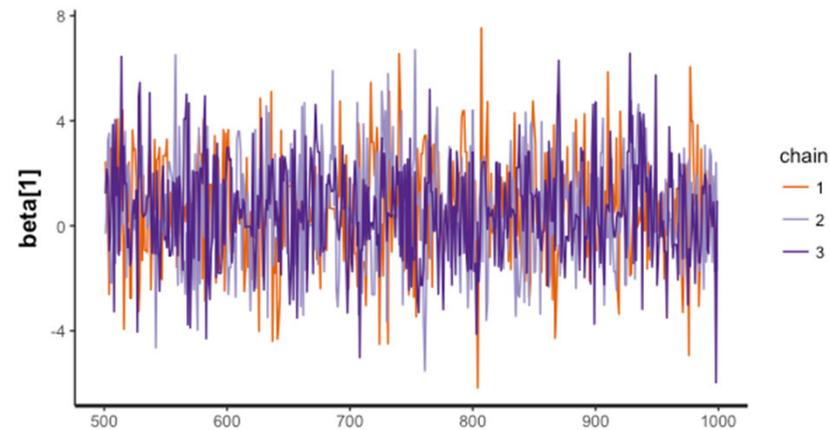
```
pars = extract(lm_intercept)
```

```
hist(pars$beta, 40, col="grey", xlab="Intercept",  
main="")
```

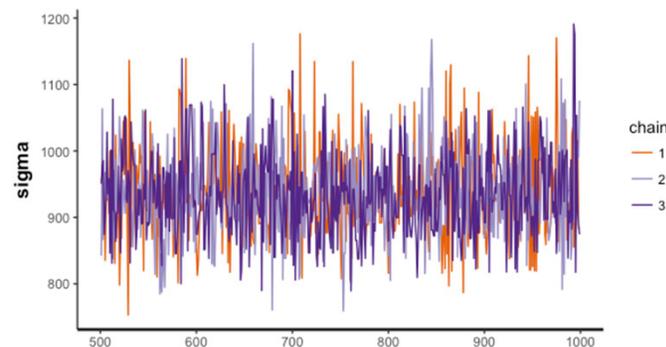
```
quantile(pars$beta, c(0.025,0.5,0.975))
```

# Traceplots of parameters

- `traceplot(lm_intercept, pars = "beta[1]")`

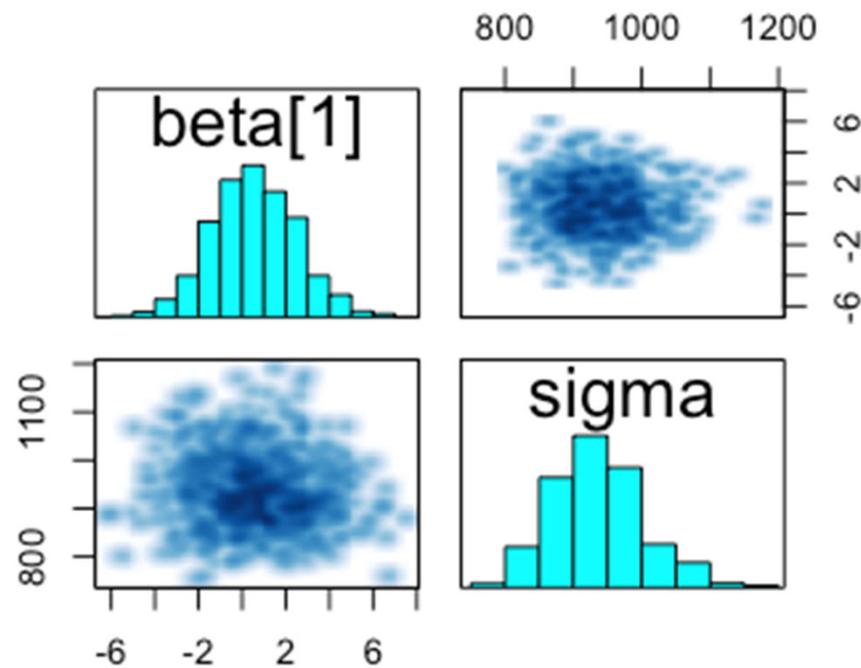


- `traceplot(lm_intercept, pars = "sigma")`



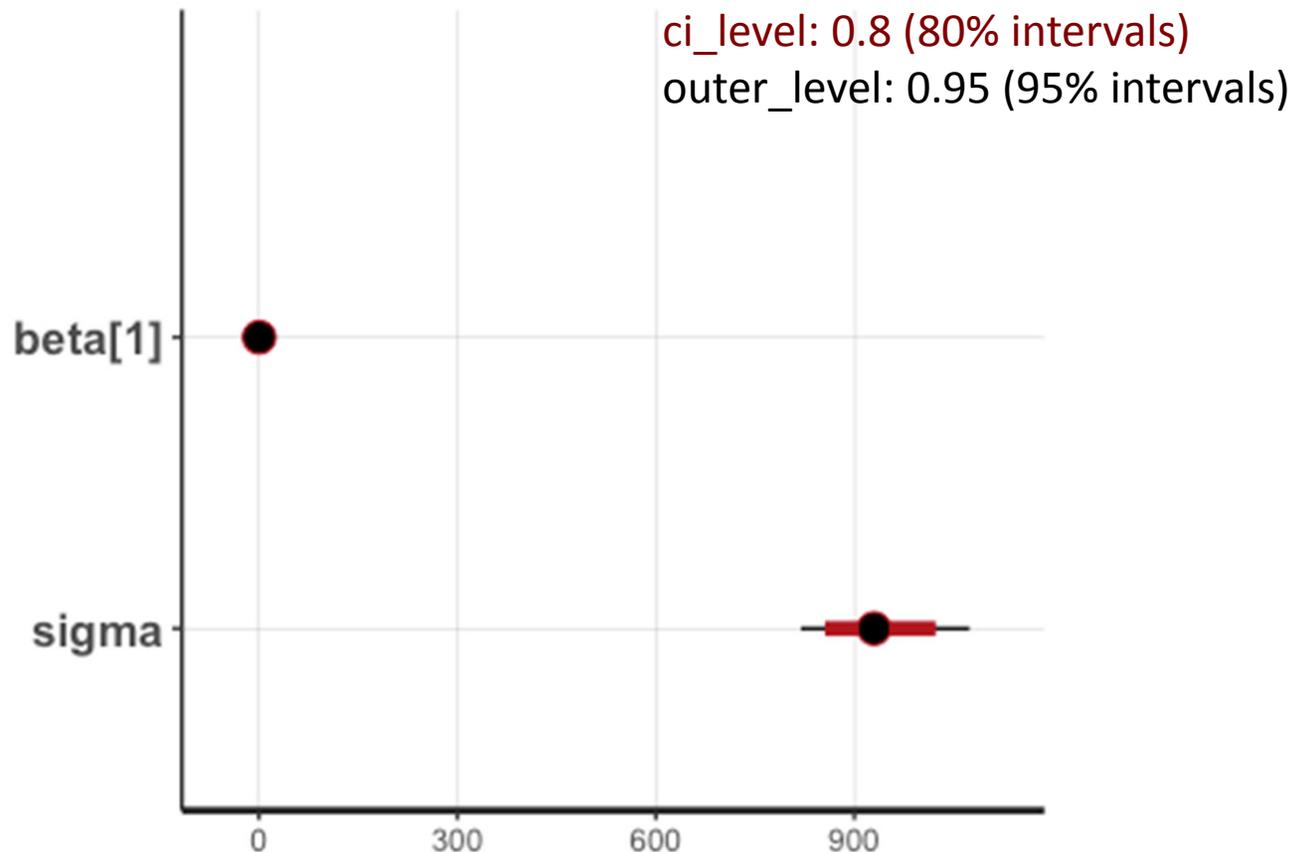
# Pairs plots between parameters

- `pairs(lm_intercept, pars=c("beta[1]","sigma"))`



# Plots with credible intervals

- `plot(lm_intercept, pars=c("beta[1]","sigma"))`



# Getting tidy summaries

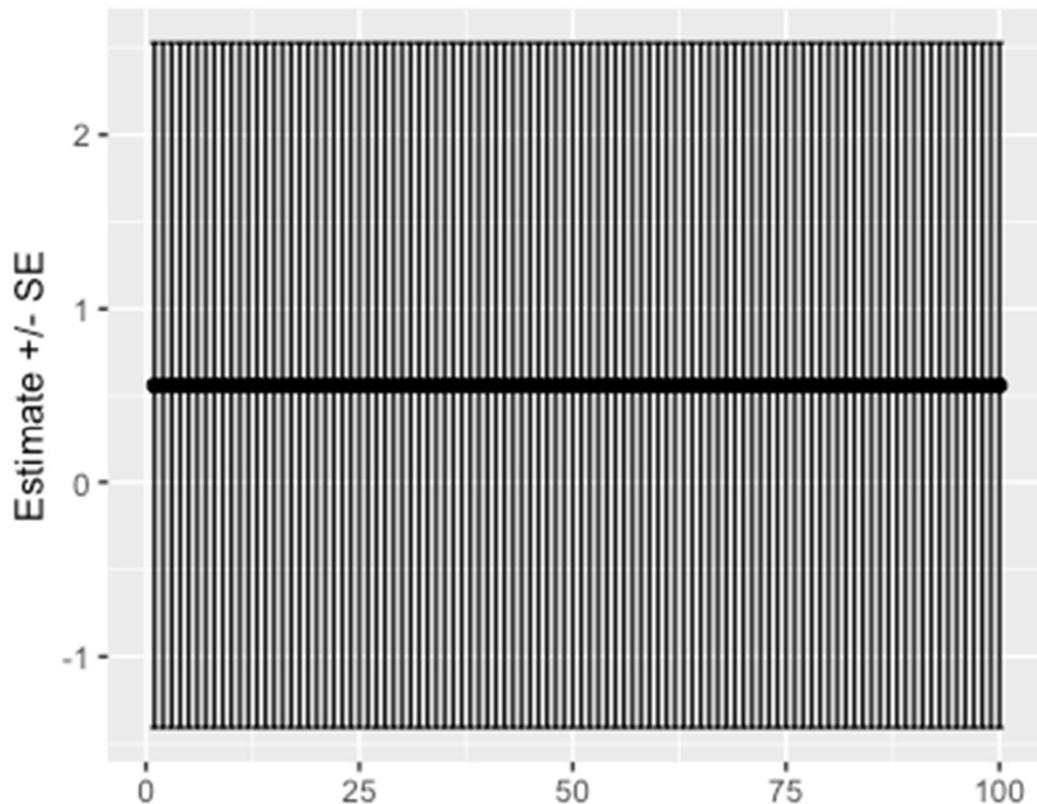
```
coef = broom::tidy(lm_intercept)
```

```
> broom::tidy(lm_intercept)
```

	term	estimate	std.error
1	beta[1]	0.5602528	1.966926
2	sigma	936.3438793	66.160572
3	pred[1]	0.5602528	1.966926
4	pred[2]	0.5602528	1.966926
5	pred[3]	0.5602528	1.966926
6	pred[4]	0.5602528	1.966926
7	pred[5]	0.5602528	1.966926
8	pred[6]	0.5602528	1.966926
9	pred[7]	0.5602528	1.966926
10	pred[8]	0.5602528	1.966926
11	pred[9]	0.5602528	1.966926
12	pred[10]	0.5602528	1.966926
13	pred[11]	0.5602528	1.966926
14	pred[12]	0.5602528	1.966926
15	pred[13]	0.5602528	1.966926
16	pred[14]	0.5602528	1.966926
17	pred[15]	0.5602528	1.966926
18	pred[16]	0.5602528	1.966926
19	pred[17]	0.5602528	1.966926

# Useful in ggplot, for example

- `ggplot(coef[grep("pred",coef$term),], aes(x = 1:100,y=estimate)) + geom_point() + ylab("Estimate +/- SE")+ geom_errorbar(aes(ymin=estimate-std.error, ymax=estimate+std.error)) + xlab("")`



Not interesting – all values are the same!

# Preserving MCMC chains

- Each chain is independent
- Defaults to merging samples from all chains together

`extract(object, pars, permuted = TRUE)`

- But summaries can be generated for each combination of parameters-chains by setting

`extract(object, pars, permuted = FALSE)`

# Random walk

- Formula:  $E[Y_t] = Y_{t-1} + e_{t-1}$
- We'll fit model to temperature data

`data(airquality)`

`Temp = airquality$Temp # air temperature`

`rw = fit_stan(y = Temp, est_drift = FALSE,  
model_name = "rw")`



# Univariate state space models

- State equation

$$x_t = \phi * x_{t-1} + e_{t-1}; e_{t-1} \sim \text{Normal}(0, q)$$

- Observation equation

$$Y_t \sim \text{Normal}(x_t, r)$$

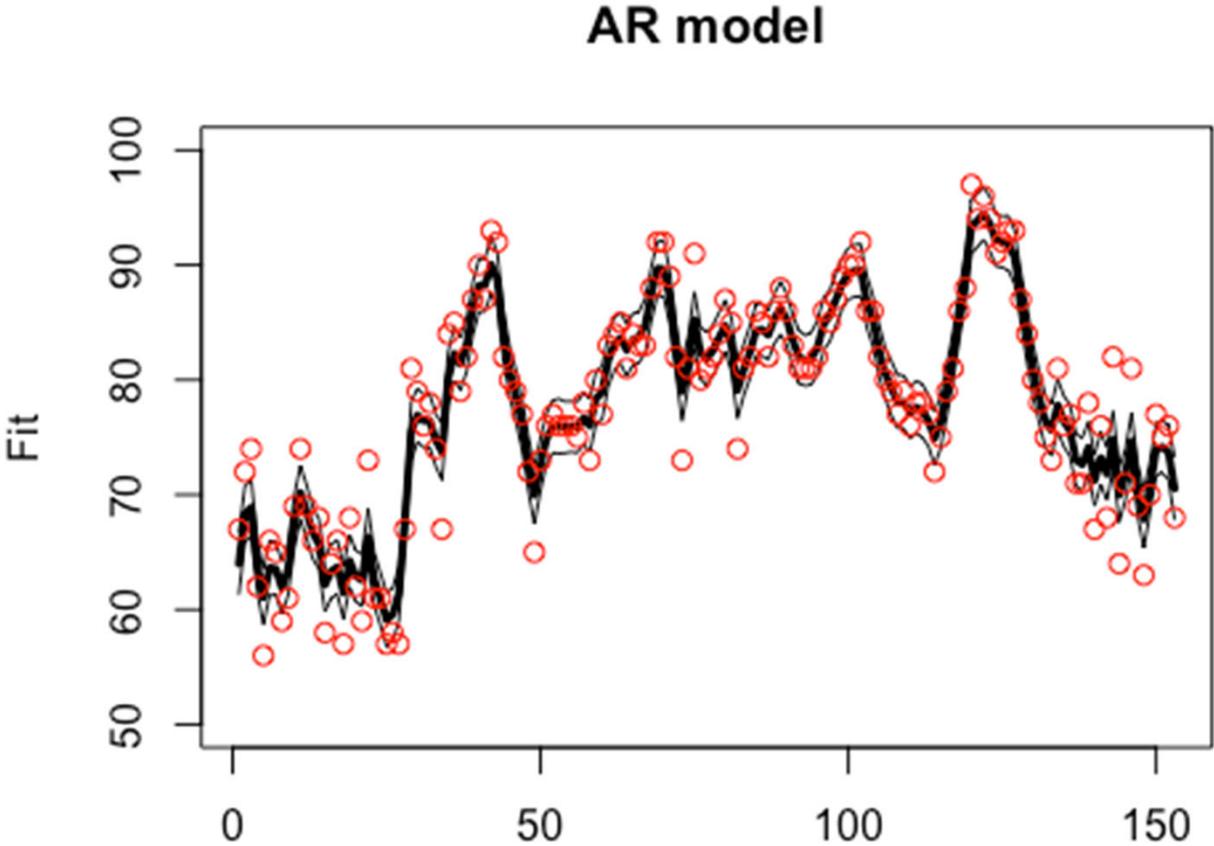
- Let's compare models with and without the ar parameter phi in the process model

# Using our stan\_fit function

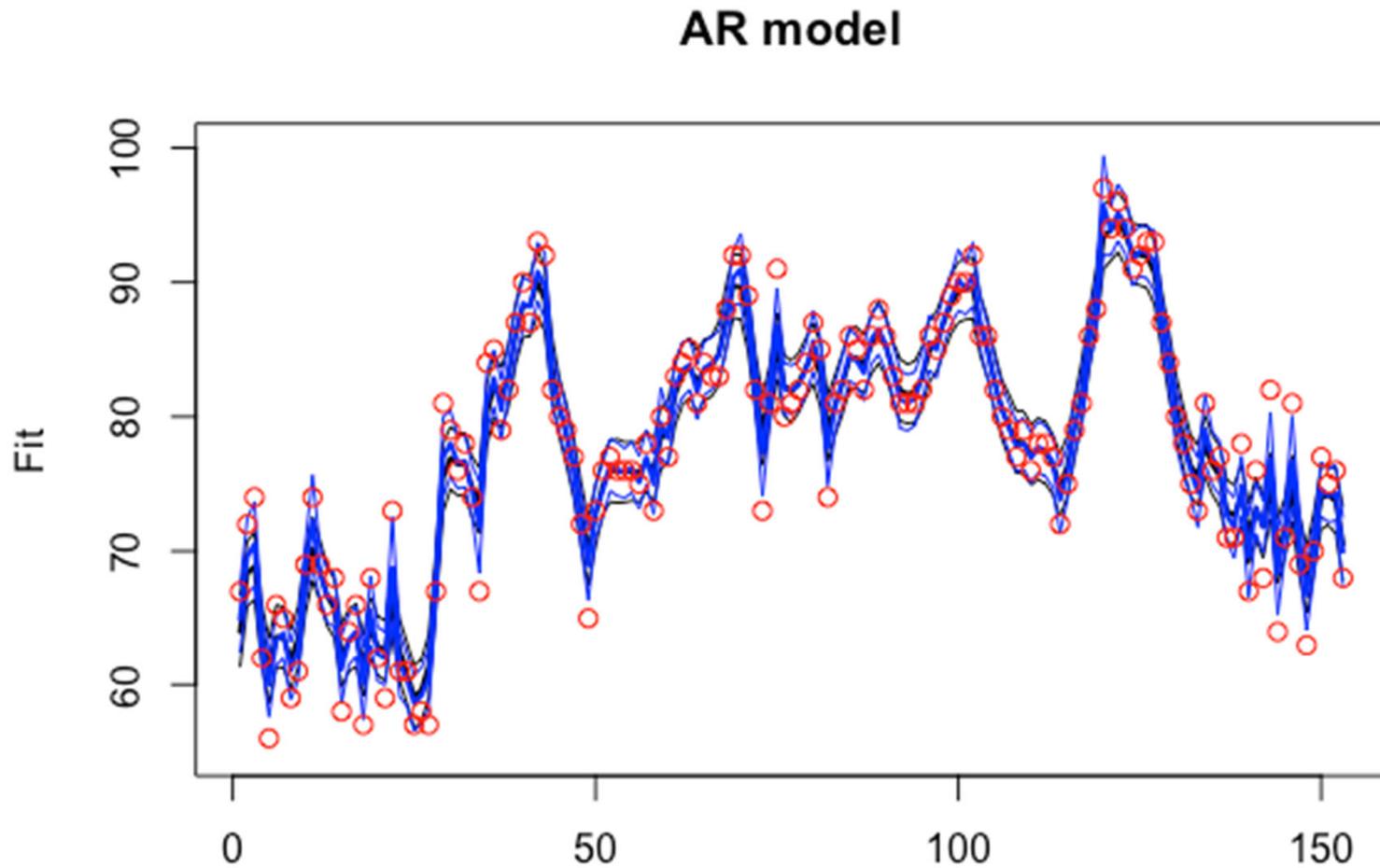
```
ss_ar = fit_stan(y = Temp, est_drift=FALSE,  
model_name = "ss_ar")
```

```
ss_rw = fit_stan(y = Temp, est_drift=FALSE,  
model_name = "ss_rw")
```

# Estimates from AR SS model

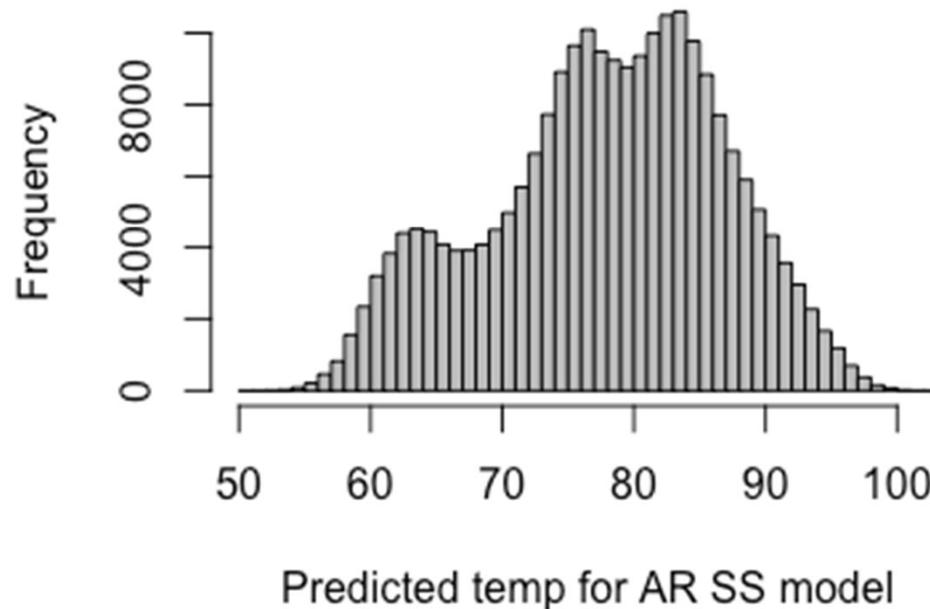


# Estimates from both models



# Posterior probability

- What's the probability that the temperature exceeds some threshold? 100 degrees?



# Probability of > 100 degrees

```
pars = extract(ss_ar)  
length(which(pars$pred > 100))
```

Low probability:  $\sim 20 / 229500$

# Dynamic Factor Analysis

- Lake WA plankton example used in manual

```
# load the data (there are 3 datasets contained here)
```

```
data(lakeWAp plankton)
```

```
# we want lakeWAp planktonTrans, which has been transformed
```

```
# so the 0s are replaced with NAs and the data z-scored
```

```
dat = lakeWAp planktonTrans
```

```
# use only the 10 years from 1980-1989
```

```
plankdat = dat[dat[,"Year"]>=1980 & dat[,"Year"]<1990,]
```

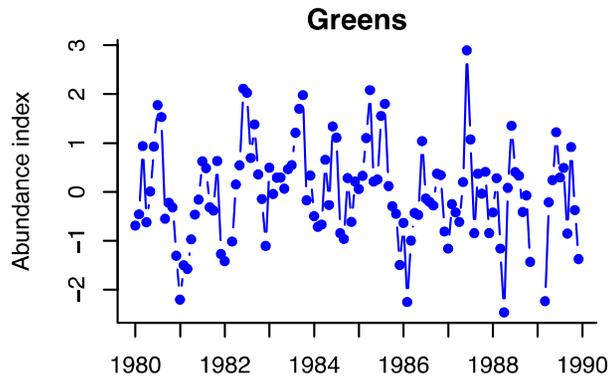
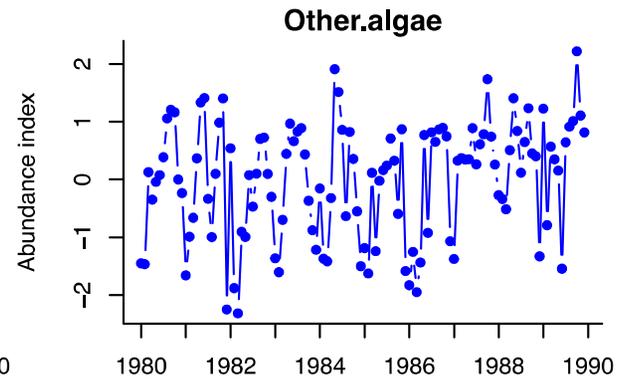
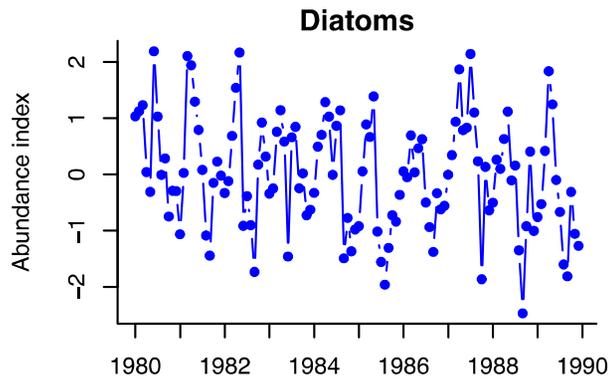
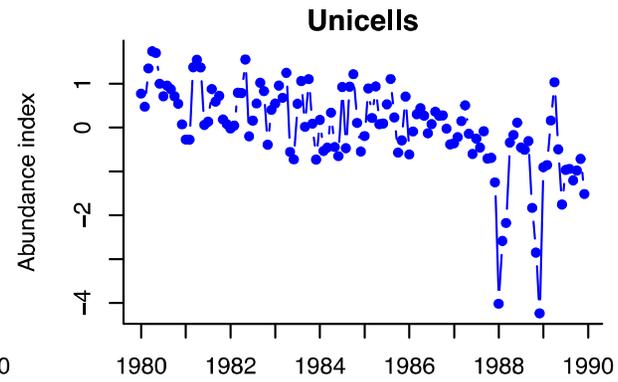
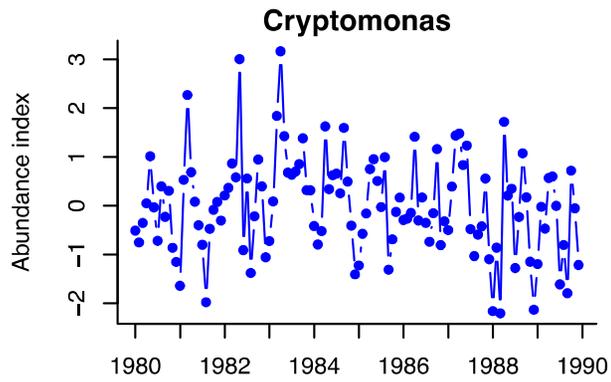
```
# create vector of phytoplankton group names
```

```
phytoplankton = c("Cryptomonas", "Diatoms", "Greens",  
                  "Unicells", "Other.algae")
```

```
# get only the phytoplankton
```

```
dat.spp.1980 = plankdat[,phytoplankton]
```

# Plankton data



# Running the model

- 3 trend model to start

```
mod_3 = fit_dfa(y = t(dat.spp.1980),  
num_trends=3)
```

# Trends need to be rotated (like MARSS)

- Again we'll use varimax rotation
- Use function we've written, **rotate\_trends**

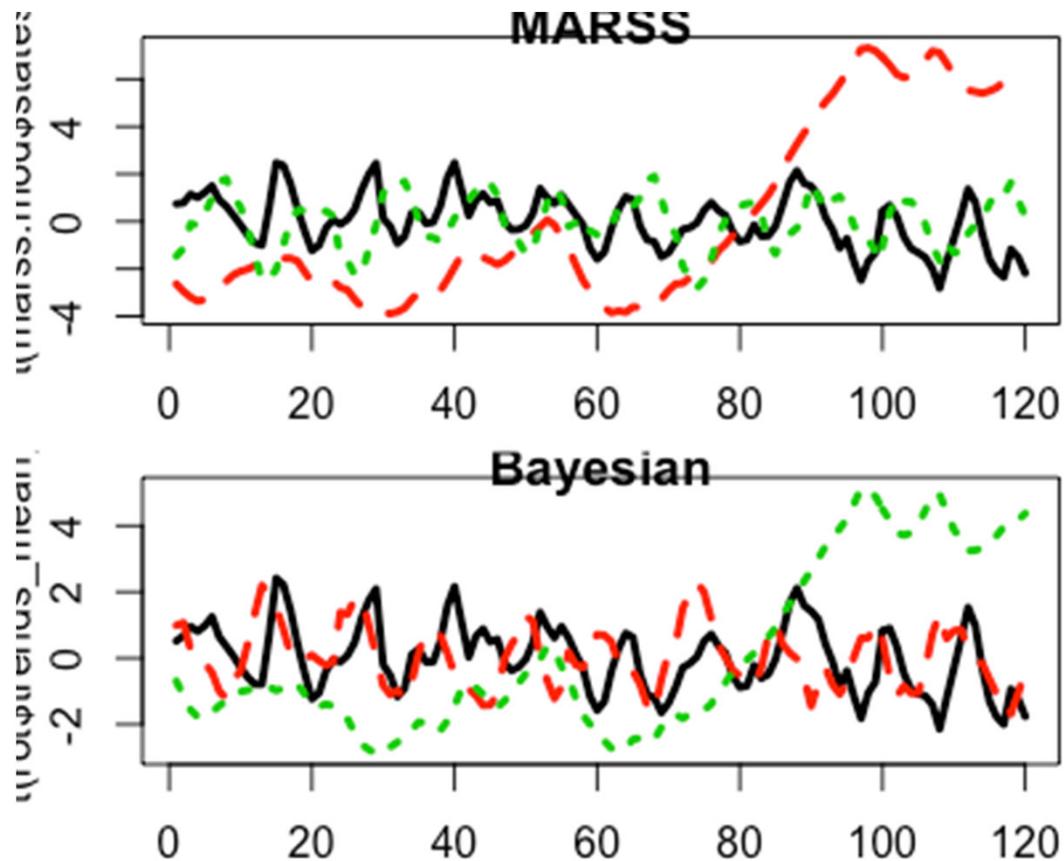
```
rot = rotate_trends(mod_3)
```

# names(rot)

- Z\_rot, rotated Z matrix for each MCMC draw
- trends, rotated trends for each MCMC draw
- **Z\_rot\_mean, mean Z across draws**
- **trends\_mean, mean trends across draws**
- trends\_lower, lower 2.5% on trends
- trends\_upper, upper 97.5% on trends

# Predicted values from Bayesian DFA

- Same results as MARSS (**trends\_mean**)



# Other variance structures

```
mod_3 = fit_dfa(y = t(dat.spp.1980),  
num_trends=3)
```

- By default, this is modeling 'diagonal and equal'
- Diagonal and unequal or shared variances can also be specified using 'varIndx' argument

```
mod_3 = fit_dfa(y = t(dat.spp.1980),  
num_trends=3, varIndx = rep(1,5))
```

# Model selection: how to select best number of trends?

- First run multiple models with varying trends

```
mod_1 = fit_dfa(y = t(dat.spp.1980), num_trends=1)
```

```
mod_2 = fit_dfa(y = t(dat.spp.1980), num_trends=2)
```

```
mod_3 = fit_dfa(y = t(dat.spp.1980), num_trends=3)
```

```
mod_4 = fit_dfa(y = t(dat.spp.1980), num_trends=4)
```

```
mod_5 = fit_dfa(y = t(dat.spp.1980), num_trends=5)
```

- 3 minutes to fit all models (4000 iterations) – probably could be at least cut in half

# Leave One Out Information Criterion (LOOIC)

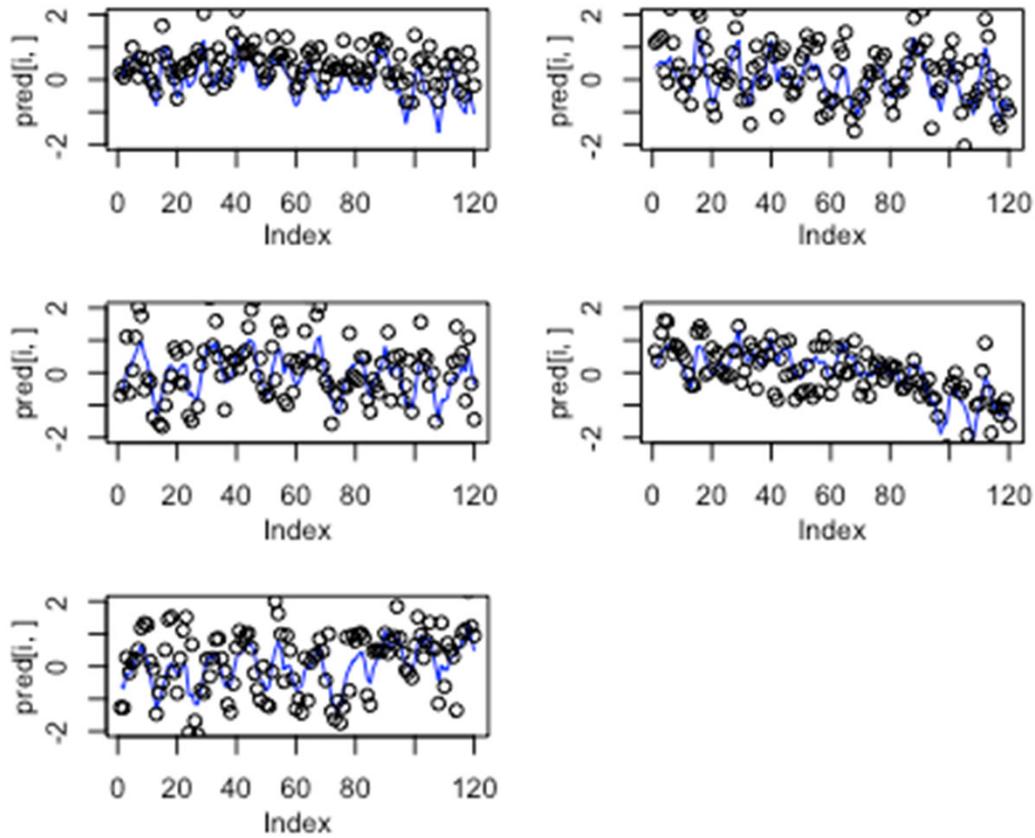
- Like AIC, lower is better
- Simple function in **library(loo)**

**loo(extract\_log\_lik(mod\_1))\$looic**

Trends	LOOIC
1	1598.889
2	1535.885
3	1469.439
4	1455.7
5	1449.304

# Predicted values

- Also like MARSS, use `$pred` parameter

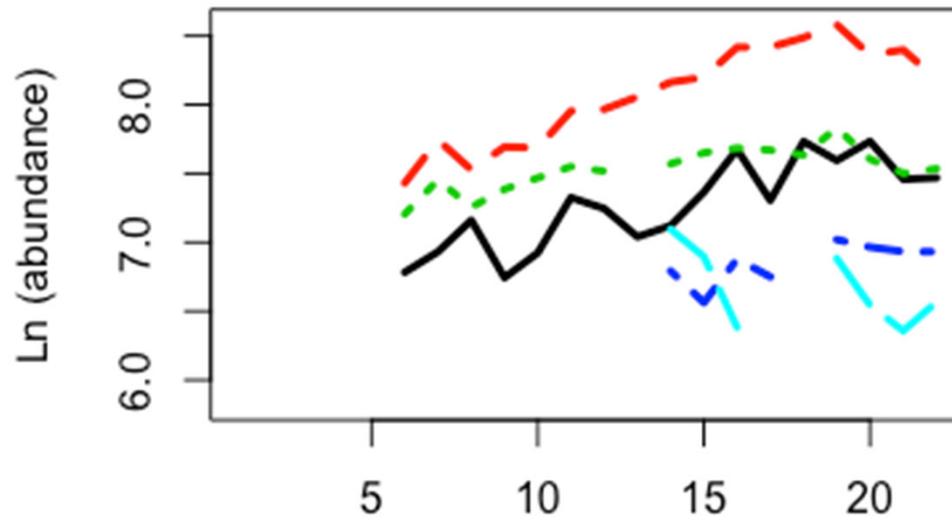


# Uncertainty intervals on states

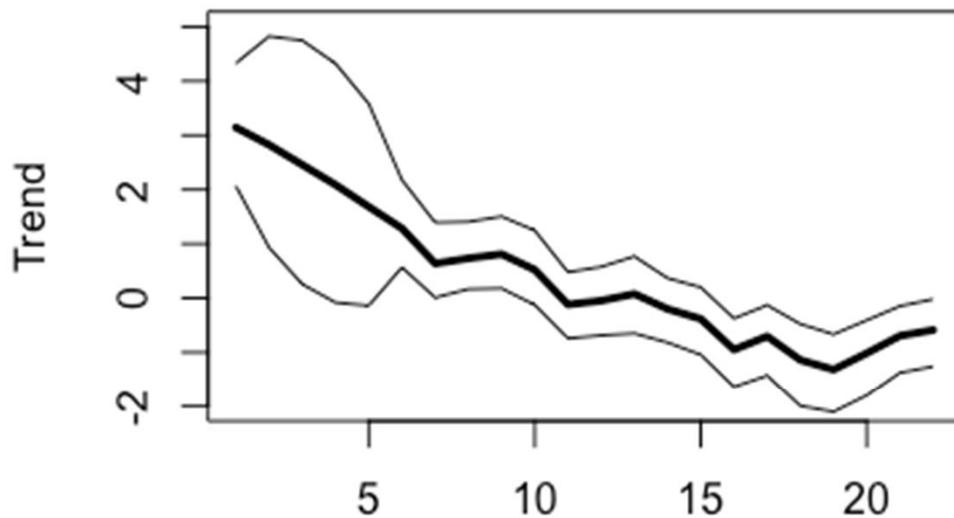
- We often just present the trend estimates for DFA – but not uncertainty
- Let's look at effect of missing data on DFA for the harbor seal dataset

```
data("harborSealWA")
```

- Assume single trend for the population

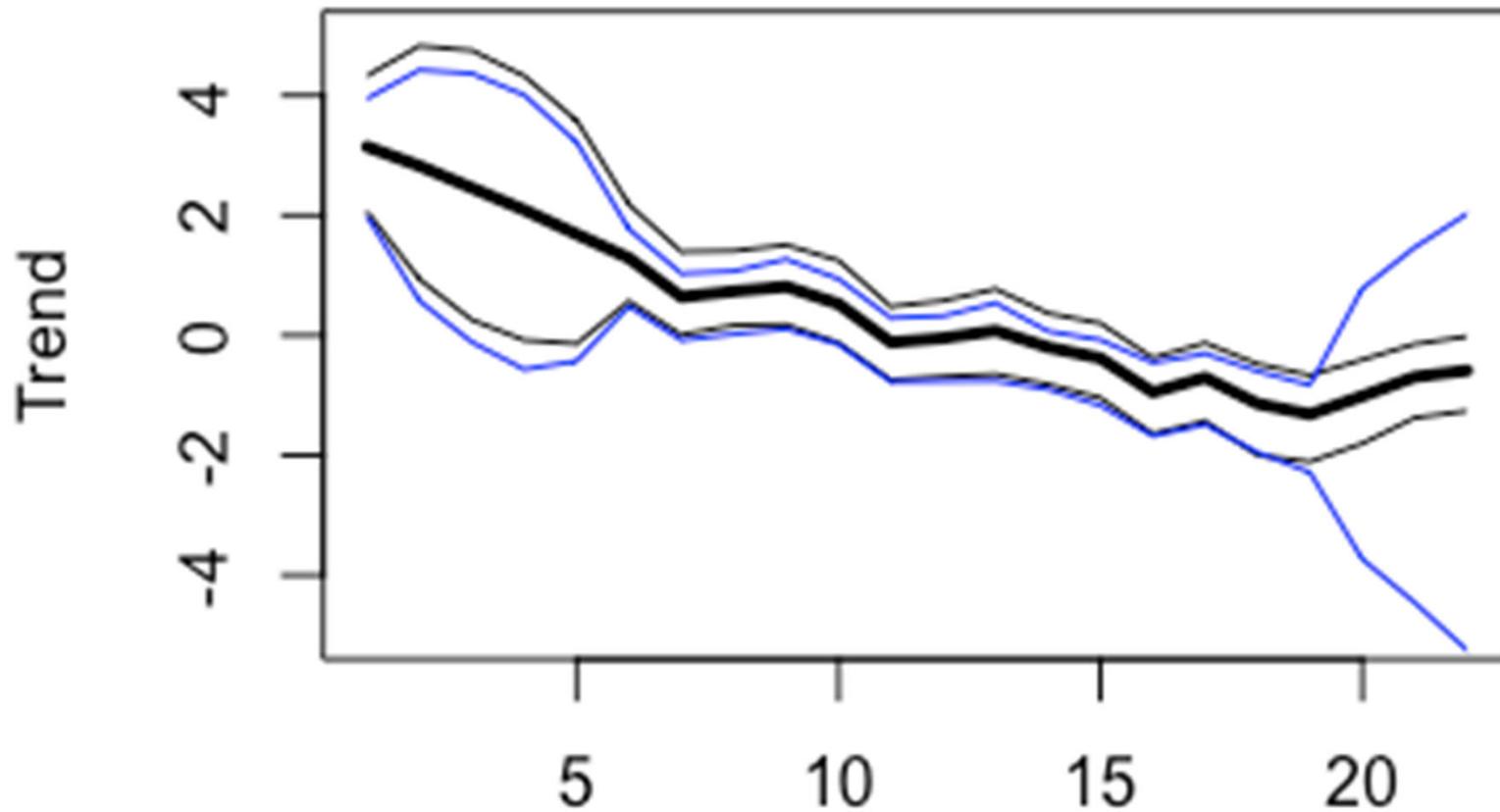


- `pars = extract(fit_dfa(y = t(harborSealWA[,-1]), num_trends = 1))`



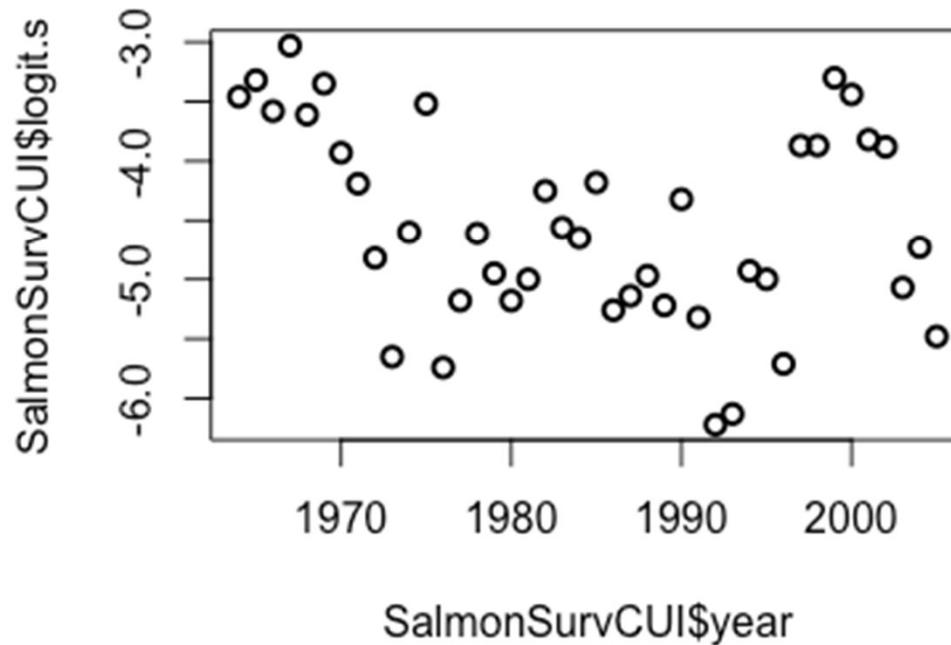
```
> harborSealWA
      Year      SJF      SJI      EBays      PSnd      HC
[1,] 1978 6.033086 6.747587 6.626718 5.820083 6.595781
[2,] 1979      NA      NA      NA      NA      NA
[3,] 1980      NA      NA      NA      NA      NA
[4,] 1981      NA      NA      NA      NA      NA
[5,] 1982      NA      NA      NA      NA      NA
[6,] 1983 6.783325 7.431300 7.205635      NA      NA
[7,] 1984 6.932448 7.744137 7.454141      NA      NA
[8,] 1985 7.160846 7.527794 7.255591 6.595781      NA
[9,] 1986 6.744059 7.693026 7.385851      NA      NA
[10,] 1987 6.923629 7.686621 7.467942      NA      NA
[11,] 1988 7.325149 7.954021 7.550661      NA      NA
[12,] 1989 7.245655 7.966933 7.516977      NA      NA
[13,] 1990 7.040536 8.057377      NA      NA      NA
[14,] 1991 7.121252 8.163371 7.569928 6.792344 7.095064
[15,] 1992 7.365180 8.199739 7.650645 6.562444 6.896694
[16,] 1993 7.675082 8.417152 7.684784 6.879356 6.383507
[17,] 1994 7.305188 8.418256 7.670429 6.749931      NA
[18,] 1995 7.732369 8.487146 7.634337      NA      NA
[19,] 1996 7.594884 8.581107 7.832411 7.020191 6.882437
[20,] 1997 7.733684 8.361007 7.604894 6.966024 6.543912
[21,] 1998 7.458186 8.398635 7.501082 6.933423 6.357842
[22,] 1999 7.468513 8.185350 7.535297 6.932448 6.566672
```

What happens when we delete last 3 years of data?



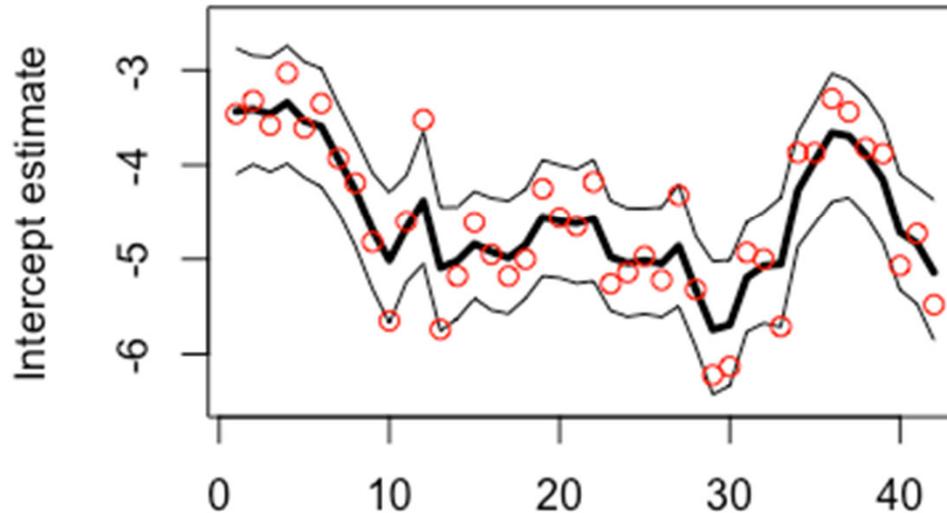
# DLMs

- Mark's example of salmon survival
- Logit transformed data



# Fitting a DLM with time varying intercept

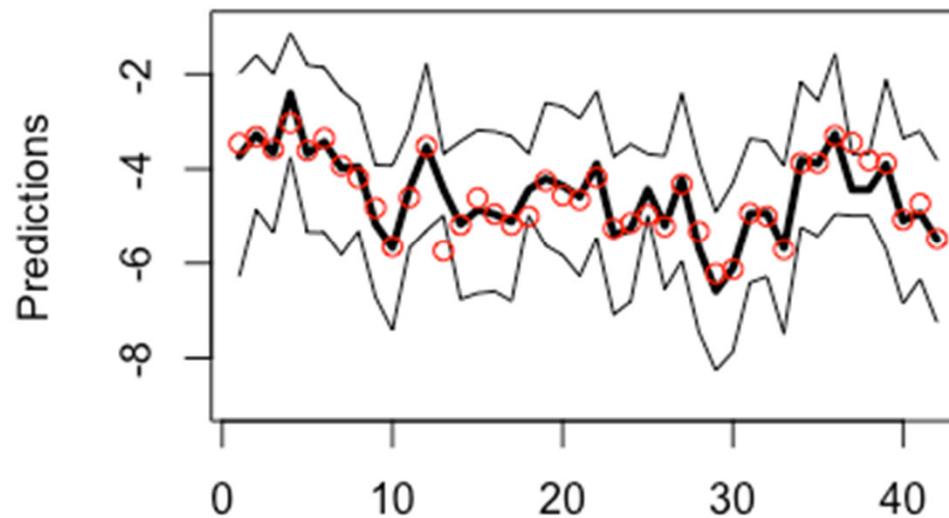
```
mod = fit_stan(y = SalmonSurvCUI$logit.s,  
model_name="dlm-intercept")
```



# Constant intercept, time – varying slope

```
mod = fit_stan(y = SalmonSurvCUI$logit.s, x =  
SalmonSurvCUI$CUI.apr, model_name="dlm-  
slope")
```

Wider CIs than  
time varying intercept  
model



# Time varying intercept and slope

- Use `model.matrix()` to specify 'x'
- `x = model.matrix(lm(SalmonSurvCUI$logit.s ~ SalmonSurvCUI$CUI.apr))`

```
mod = fit_stan(y = SalmonSurvCUI$logit.s, x =  
model.matrix(lm(SalmonSurvCUI$logit.s ~  
SalmonSurvCUI$CUI.apr)), model_name="dlm")
```

# Summary

- Additional models available, e.g. 'marss'
- Very flexible
- Easy to add custom features on existing code